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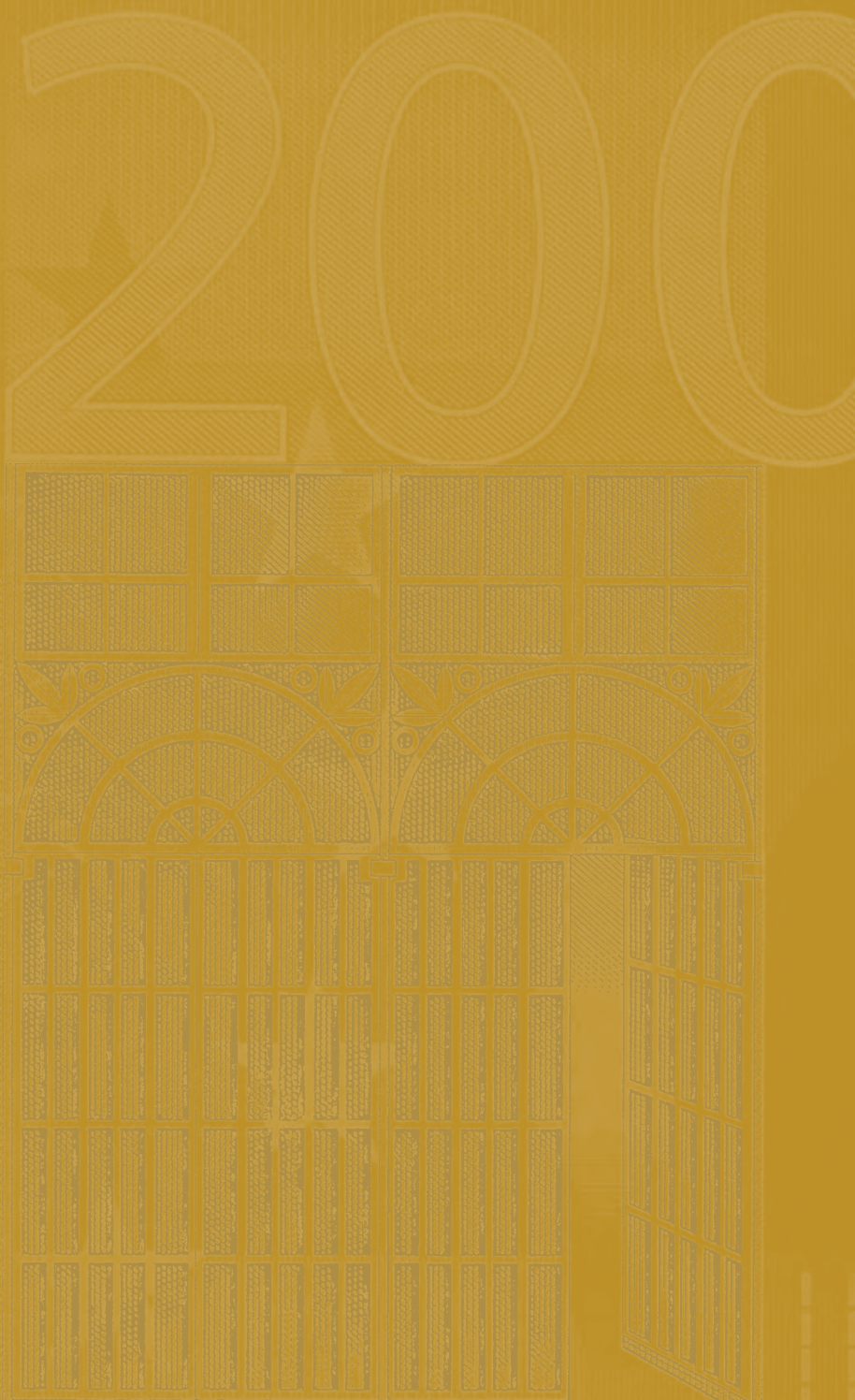
WAGE DYNAMICS
NETWORK

WORKING PAPER SERIES

NO 1053 / MAY 2009

**INFLATION DYNAMICS
WITH LABOUR
MARKET MATCHING
ASSESSING ALTERNATIVE
SPECIFICATIONS**

by Kai Christoffel, James Costain,
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by Kai Christoffel², James Costain³,
Gregory de Walque⁴, Keith Kuester⁵,
Tobias Linzert², Stephen Millard⁶
and Olivier Pierrard⁷

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² European Central Bank, Kaiserstrasse 29, D-60311 Frankfurt am Main, Germany; e-mail: kai.christoffel@ecb.europa.eu; tobias.linzert@ecb.europa.eu

³ Banco de España, Alcalá 50, E-28014 Madrid, Spain; e-mail: james.costain@bde.es

⁴ Banque Nationale de Belgique, Boulevard de Berlaimont 14, B-1000 Brussels, Belgium; e-mail: gregory.dewalque@nbb.be

⁵ Federal Reserve Bank of Philadelphia, Ten Independence Mall, Philadelphia, PA 19106-1574, USA; e-mail: keith.kuester@phil.frb.org

⁶ Bank of England, Threadneedle Street, London EC2R 8AH, UK
e-mail: stephen.millard@bankofengland.co.uk

⁷ Banque Centrale du Luxembourg, 2 boulevard Royal, L-2983 Luxembourg, Luxembourg;
e-mail: Olivier.Pierrard@bcl.lu

Wage Dynamics Network

This paper contains research conducted within the Wage Dynamics Network (WDN). The WDN is a research network consisting of economists from the European Central Bank (ECB) and the national central banks (NCBs) of the EU countries. The WDN aims at studying in depth the features and sources of wage and labour cost dynamics and their implications for monetary policy. The specific objectives of the network are: i) identifying the sources and features of wage and labour cost dynamics that are most relevant for monetary policy and ii) clarifying the relationship between wages, labour costs and prices both at the firm and macro-economic level.

The WDN is chaired by Frank Smets (ECB). Giuseppe Bertola (Università di Torino) and Julian Messina (Universitat de Girona) act as external consultants and Ana Lamo (ECB) as Secretary.

The refereeing process of this paper has been co-ordinated by a team composed of Gabriel Fagan (ECB, chairperson), Philip Vermeulen (ECB), Giuseppe Bertola, Julian Messina, Jan Babecký (CNB), Hervé Le Bihan (Banque de France) and Thomas Mathä (Banque centrale du Luxembourg).

The paper is released in order to make the results of WDN research generally available, in preliminary form, to encourage comments and suggestions prior to final publication. The views expressed in the paper are the author's own and do not necessarily reflect those of the ESCB.

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Address

Kaiserstrasse 29
60311 Frankfurt am Main, Germany

Postal address

Postfach 16 03 19
60066 Frankfurt am Main, Germany

Telephone

+49 69 1344 0

Website

<http://www.ecb.europa.eu>

Fax

+49 69 1344 6000

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Abstract

This paper reviews recent approaches to modeling the labour market and assesses their implications for inflation dynamics through both their effect on marginal cost and on price-setting behaviour. In a search and matching environment, we consider the following modeling setups: right-to-manage bargaining *vs.* efficient bargaining, wage stickiness in new and existing matches, interactions at the firm level between price and wage-setting, alternative forms of hiring frictions, search on-the-job and endogenous job separation. We find that most specifications imply too little real rigidity and, so, too volatile inflation. Models with wage stickiness and right-to-manage bargaining or with firm-specific labour emerge as the most promising candidates.

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Non-Technical Summary

The main task of central banks is to maintain price stability by controlling inflation and, for this reason, it is important to understand what drives the dynamics of inflation. A long tradition in monetary economics has assigned the labour market a central role in inflation dynamics. Not least given the wide diversity in labour market structures, institutions and policies across Europe, and the world more generally, it is of interest to know whether or not heterogeneity in certain aspects of the labour market matters for inflation dynamics and, hence, monetary policy.

This paper analyses and compares existing approaches to modelling the labour market, explaining their implications for the behaviour of inflation. The paper highlights which of the particular features of each modelling approach are most important for driving inflation dynamics and provides a structure to the rich variety of modelling approaches in previous work. In so doing, we follow an active strand of research that has set out explicitly to model individual workers and firms, who take time to form job matches. The rationale for so doing comes from the belief that the slow responses of employment and unemployment to changes in demand are a natural place to look for the origins of the slow response of inflation to changes in demand.

We start by examining what euro-area data suggests happens to labour market variables after an unexpected fall in interest rates (ie, a monetary policy shock). We find that output rises significantly above its steady state, that inflation rises, that wages per employee also rise but by less than output (in percentage terms), that employment rises significantly, and unemployment falls and, finally, that most of the adjustment in labour is borne by the number of employees rather than by hours worked per employee.

We then consider a standard macroeconomic model with unemployed workers looking for jobs, firms posting vacancies, and a clear distinction between employment and hours worked. Building on this, we replace certain assumptions about the labour market structure by others, one at a time. First, we consider alternative ways that firms and workers can decide on how many hours each worker works. Then, we consider the effect of wage negotiations being staggered, with wages fixed for some time and what happens when we, additionally, tie the wages of newly hired workers to those of existing workers in a firm. Next, we consider interactions at the firm-level between price and wage-setting. We then consider various types of hiring costs before moving on to consider search on-the-job, and finish by considering what happens when job destruction varies over the cycle. In each of the cases we provide intuition for the effect that a specific modification of the baseline model has on inflation dynamics.

We find that the baseline model predicts a response of inflation to changes in interest rates that is too large relative to the data. Allowing search on the job and considering different types of hiring cost does not seem to affect this result. However, when our baseline model is

combined with the assumption that once employers and employees have agreed on an hourly wage, the employer chooses how many hours his employees will work, staggered wages help to smooth the reaction of wages resulting in a smaller response of inflation to an interest rate changes. Inflation responds even less when we account for the firm-specific nature of labour. But, in this case, the model also has implications for the responses of unemployment and vacancies that do not match the data.

More generally, by analysing a wide range of institutional features of the labour market, we show that only those institutional features that affect or generate a direct channel from wages to inflation matter for how inflation responds to interest rate changes. By contrast, institutional features that leave this channel unaffected matter much less for inflation dynamics.

1 Introduction and motivation

The key task of central banks is to maintain price stability by controlling inflation, and, for this reason, it is important to understand what drives the dynamics of inflation. A crucial element of this is the underlying nature of nominal and real frictions associated with the adjustments of prices in the economy. A long tradition in monetary economics, starting with Phillips (1958), has assigned labour market frictions and, in particular wage-setting frictions, a central role in inflation dynamics. Not least given the wide diversity in labour market structures, institutions, and policies across Europe, and the world more generally, it is of interest to know whether heterogeneity in certain aspects of the labour markets matters for inflation dynamics and, hence, monetary policy.

This paper analyses and compares existing labour market modeling approaches within the New Keynesian business cycle model setup. In particular, it explains their implications for the behaviour of marginal cost and inflation in theory and quantitatively using a calibration with common parameters based on euro-area data. The paper highlights which of the particular features of each modeling approach are key for driving inflation dynamics and provides a structure to the rich variety of modeling approaches in the existing literature.¹ We follow an active strand of research that has set out to model the labour market more explicitly within the New Keynesian business cycle model, most often opting for including variants of the Mortensen and Pissarides (1994) model of search and matching frictions. The rationale for so doing comes from the belief that sluggish responses in labour market variables to shocks are a natural place to look for the origins of the sluggish response of inflation to shocks. In terms of the New Keynesian framework, labour market frictions will alter either aggregate marginal cost or the firms' price-setting behaviour for a given marginal cost.² However, so far, the implications for inflation dynamics seem to be mixed. Some have, for example, argued that accounting for equilibrium unemployment increases the resilience of marginal cost, and hence inflation, to shocks by adding a pool of slack resources; see, e.g., Trigari (2009) and Walsh (2005). Others, e.g., Krause and Lubik (2007), conclude that contrary to received wisdom wage rigidity does not have an effect on inflation inertia.³ As a starting point, we choose the efficient bargaining version of Trigari (2006): that

¹Other papers, largely sparked by Shimer's (2005) critique, seek to improve the modeling or calibration of the labour market *per se* and are not necessarily concerned with nominal frictions, e.g., Gertler and Trigari (2006), Fujita and Ramey (2007), and Yashiv (2006). In this paper, we take up some of the suggestions made in these papers, since – through their implications for the behaviour of wages and unemployment – they also imply changes in the behaviour of real marginal cost.

²Thus, the labour market was seen as a source of 'real rigidities'. For an overview of the extensive literature on real rigidities more generally, see Woodford (2003).

³Another strand of literature has been concerned with normative implications, such as the design of optimal monetary policy; see, e.g., Blanchard and Gali (2006), Faia (2008), Tang (2006).

is, a plain New Keynesian model with search and matching frictions and a distinction between heads and hours. Building on this, we replace certain assumptions on the labour market structure by others, one at a time. We consider the following variants: right-to-manage (instead of efficient) bargaining (Trigari (2006)), stickiness of wages in these setups (Gertler and Trigari (2006) and Christoffel and Linzert (2005)) differentiating between stickiness in new matches and existing matches (Bodart *et al.* (2005) and Bodart *et al.* (2006)), interactions at the firm-level between price and wage-setting (Kuester (2007), Sveen and Weinke (2007) and Thomas (2008)), alternative forms of vacancy posting costs (Yashiv (2006)) and the hiring process (Fujita and Ramey (2007)), search on-the-job (Krause and Lubik (2006) and van Zandweghe (2009)) and endogenous separation (den Haan *et al.* (2000) and Zanetti (2007)). In each of the cases we provide intuition for the effect that a specific modification of the baseline model has on inflation dynamics.

Our findings are as follows. We show, first, that the baseline model predicts a response of inflation that is too large relative to the data, as a result of the large and immediate response of real marginal cost. This would suggest either that the labour market is not an important source of real rigidities or that the specific way of modeling the labour market does not capture the way the labour market affects inflation dynamics in practice. We follow the second route. When combined with a ‘right-to-manage’ assumption for the determination of hours, staggered wages at the level of the match help to smooth the reaction of the marginal wage resulting, in a smaller response to shocks of marginal cost and inflation. Inflation responds even less to shocks when we account for the firm-specific nature of labour in the search and matching model. Doing so, however, also has implications for the responses of unemployment and vacancies that take these further away from the data. Variations of the hiring costs have only a small impact on inflation dynamics. In contrast, we find that endogenous separation may be a promising candidate for reducing the response of inflation to shocks, though this finding depends critically on the calibration, as is the case for many of the examples. More generally, by analysing a wide range of institutional features of the labour market employing a variety of variants of the Mortensen and Pissarides (1994) search and matching model, we show that only those institutional features that affect or generate a direct channel from wages to inflation matter for inflation dynamics. By contrast, institutional features that leave this channel unaffected matter much less for inflation dynamics.

The rest of the paper is structured as follows. In section 2, we lay out the baseline model. We highlight which features of the model will be changed subsequently. In section 3, we calibrate the baseline model to euro area data. Subsequently, we compare

the responses of inflation (and wages and unemployment) to a monetary policy shock in the model with the responses to the same shock in euro area data. In section 4 we consider the effects of having ‘right-to-manage’ bargaining and section 5 adds nominal wage rigidities for new and existing jobs, exploring how these rigidities interact with the bargaining scheme. Section 6 considers what happens if wages and prices are set simultaneously in the presence of firm-specific labour. Section 7 allows for new hires being productive immediately and not only with a one-period delay. In section 8 we discuss the effects of varying the free-entry condition and vacancy costs. Sections 9 and 10 consider other margins along which adjustment can occur in the labour market: section 9 discusses on-the-job search and section 10 discusses endogenous job destruction. Section 11 concludes. An Appendix, as a point of reference, presents empirical evidence on monetary transmission in the euro area.

2 The baseline model

In this section we lay out our baseline model, a simplified version of Trigari’s (2006) efficient bargaining model.

2.1 Households, consumption and saving

We assume that there is a continuum of workers indexed by j on the unit interval who supply a homogeneous type of labour. Only a proportion n_t of them are employed. We adopt a representative – or large – household interpretation so that the unemployment rate $u_t = 1 - n_t$ is identical at the household or at the aggregate level. As shown by Merz (1995) the representative household assumption amounts to allowing for the existence of state-contingent securities offering workers insurance against differences in their specific labour income. Household members (workers) share their labour income, i.e., wage and unemployment benefits, before choosing per capita consumption and bond holdings. Labour income is made of the sum of hourly wages, w_t , weighted by total hours worked, $n_t h_t$, and unemployment benefits, b , weighted by the number of unemployed, $1 - n_t$.⁴ As shareholders, households also receive the profits Ψ_t generated by the monopolistic competitive retail firms and the intermediate producers of labour services.

We assume separability between leisure and consumption in the instantaneous utility function, implying that all workers share the same marginal utility of wealth and choose the same optimal consumption, be they employed or not. A worker’s utility can be

⁴ b could alternatively be interpreted as the income generated by the domestic activities of an unemployed worker.

written as

$$\mathcal{U}(c_{jt}, h_{jt}) = \frac{c_{jt}^{1-\sigma_c}}{1-\sigma_c} - \chi \frac{h_{jt}^{1+\phi}}{1+\phi} \quad (1)$$

with $\sigma_c \geq 1, \chi > 0, \phi \geq 0$. Let $\mathcal{H}_t(j)$ be the value function of worker j . If we momentarily leave aside the labour supply decision, which is taken by the household as a whole, worker j 's maximisation program is

$$\mathcal{H}_t(j) = \max_{c_{jt}, B_{jt+1}} \{ \mathcal{U}(c_{jt}, h_{jt}) + \beta \cdot \mathbb{E}_t \mathcal{H}_{t+1}(j) \}$$

subject to

$$\frac{B_t}{r_t p_t} + c_t + t_t = \frac{B_{t-1}}{p_t} + n_t w_t h_t + (1 - n_t) b + \Psi_t \quad (2)$$

where workers hold their financial wealth in the form of one-period bonds, B_t , with price $1/r_t$ (where r_t is the nominal interest rate), c_t represents aggregate consumption, t_t are lump-sum taxes payable and p_t is the consumer price index.

The worker's optimal consumption and saving decision coincide with those of his peers and are derived from the following first-order conditions:

$$\frac{\lambda_t}{p_t} = \beta r_t \mathbb{E}_t \left\{ \frac{\lambda_{t+1}}{p_{t+1}} \right\} \quad (3)$$

$$\lambda_t = c_t^{-\sigma_c} \quad (4)$$

2.2 Firms

In our baseline model, the economy consists of three sectors: a perfectly competitive final good sector, a monopolistically competitive retail sector and a perfectly-competitive intermediate good sector.

2.2.1 Final goods sector

We assume a continuum of differentiated goods indexed by i on the unit interval. Final good firms aggregate the differentiated goods $y_t(i)$ produced by the retailers into a homogeneous consumption good using a Dixit-Stiglitz technology

$$y_t = \left[\int_0^1 [y_t(i)]^{\mu_p} di \right]^{1/\mu_p} \quad \text{with } \mu_p \in (0, 1) \quad (5)$$

where $1/\mu_p$ represents the retailers' gross price mark-up while $1/(1 - \mu_p)$ is the elasticity of substitution between intermediate differentiated goods. Each final good firm

maximises profit, leading to the following demand for intermediate good i

$$y_t(i) = \left(\frac{p_t(i)}{p_t} \right)^{\frac{1}{\mu_p - 1}} y_t \quad (6)$$

where p_t is the final good price, obtained by aggregation of the retailers' prices

$$p_t = \left[\int_0^1 [p_t(i)]^{\frac{\mu_p}{\mu_p - 1}} di \right]^{\frac{\mu_p - 1}{\mu_p}} \quad (7)$$

The modeling of the final good sector will remain unchanged throughout the paper.

2.2.2 Monopolistic retailers and price setting

Given the demand $y_t(i)$ retail firm i faces for its product, it buys a homogeneous intermediate labour good at nominal price $p_t x_t$ per unit and transforms it one for one into a differentiated product. In each period, a fraction $1 - \xi_p$ of retail firms sets a new price $p_t^*(i)$. This price prevails j periods later with probability ξ_p^j . The price-setting firms maximise the discounted flows of expected real profits using a discount rate consistent with that used by their shareholders (i.e., households). All the price-setting firms face the same optimisation problem, implying that they all choose the same new price p_t^* . Profit maximisation results in the following first-order condition

$$\mathbb{E}_t \sum_{j=0}^{\infty} \beta_{t,t+j} (\xi_p)^j \frac{y_{t+j}(i)}{p_{t+j}} \left(p_t^* - \frac{1}{\mu_p} p_{t+j} x_{t+j} \right) = 0, \quad (8)$$

where $\beta_{t,t+j} := \beta^j \frac{\lambda_{t+j}}{\lambda_t}$ denotes the stochastic discount factor. Log-linearising the resulting expression around the steady state enables us to derive the New Keynesian Phillips curve

$$\hat{\pi}_t = \beta \mathbb{E}_t \hat{\pi}_{t+1} + \frac{(1 - \xi_p)(1 - \beta \xi_p)}{\xi_p} \hat{x}_t, \quad (9)$$

where $\pi_t = p_t/p_{t-1}$ is the inflation rate while 'hats' denote percentage deviations from steady state. This expression makes clear that, for a given response of inflation to movements in real marginal cost – which will depend on ξ_p – the response of inflation to a given shock will depend on how real marginal cost, \hat{x}_t , responds to that shock.

2.2.3 Intermediate goods sector

In the baseline model, there are a continuum of intermediate labour good producers uniformly distributed and selling their output in a competitive market. Their only



factor of production is labour, and labour efficiency is decreasing with hours, so that h hours supplied by one worker produce only h^α units of efficient labour, $\alpha < 1$. Following Mortensen and Pissarides (1999), we assume that intermediate producers can hire at most one worker so that their production is either zero or

$$y_t^i(o) = [h_t(o)]^\alpha \text{ with } \alpha < 1 ,$$

where o indicates match o . In the baseline the intermediate goods sector and the retailers are linked through competitive markets for the intermediate good. In section 6, we assess the implications for inflation dynamics once these two sectors are merged so that there are interactions between wage and price setting at the level of the individual firm.

2.3 The labour market

The labour market is organised such that it links the intermediate good firms and the workers.

2.3.1 Labour market flows

Let m_t denote the number of new firm-worker matches. We assume that the number of matches is a function of the number of job vacancies v_t and effective job seekers u_t , and we consider the following linear homogeneous matching function:

$$m_t = \sigma_m u_t^\vartheta v_t^{1-\vartheta} \tag{10}$$

with $\sigma_m > 0$ and $\vartheta \in (0, 1)$.⁵ In the baseline, the probability an unemployed worker finds a job is given by

$$s_t = \frac{m_t}{u_t} \tag{11}$$

while the probability that a firm fills a vacancy is given by

$$q_t = \frac{m_t}{v_t} . \tag{12}$$

An exogenous proportion ρ of firm-worker relationships ends each period, which implies the following employment dynamics:

$$n_t = (1 - \rho)n_{t-1} + m_{t-1} . \tag{13}$$

⁵Models with search on-the-job (section 9) modify the matching function to depend also on the number of employed workers seeking jobs.

In the baseline therefore employment is predetermined, while hours per worker are free to adjust contemporaneously. This means that marginal costs are in the first period after a shock influenced exclusively by the marginal cost of an additional hour worked on the intensive margin. We ease this assumption in section 7, where we allow for contemporaneous hiring, and in section 10, which looks at endogenous separation.

2.3.2 *Value of a job and of a vacancy for an intermediate labour-services producer*

Firms and workers bargain over wages; such bargaining distributes the rents that arise once a match is formed to both the firm and the worker. We denote by J_t and V_t , respectively, the asset value of a job and of a vacancy at period t , dropping match index i for convenience:

$$J_t = y_t^l x_t - h_t w_t + \mathbb{E}_t \{ \beta_{t,t+1} (1 - \rho) J_{t+1} \} , \quad (14)$$

$$V_t = -k_t^v + \mathbb{E}_t \{ \beta_{t,t+1} [q_t J_{t+1} + (1 - q_t) V_{t+1}] \} , \quad (15)$$

where k_t^v represents the per-period hiring cost (in units of the consumption good).

2.3.3 *Value of a job and value of unemployment for a household*

For a worker's family, which bargains on his behalf, the value of a new job is given by

$$W_t = h_t w_t - \frac{\chi}{\lambda_t} \frac{h_t^{1+\phi}}{1+\phi} + \mathbb{E}_t \{ \beta_{t,t+1} [(1 - \rho) W_{t+1} + \rho U_{t+1}] \} , \quad (16)$$

where U_t represents the present value of being unemployed at period t . Formally,

$$U_t = b + \mathbb{E}_t \{ \beta_{t,t+1} [s_t W_{t+1} + (1 - s_t) U_{t+1}] \} . \quad (17)$$

2.4 Wage bargaining

In this section, we discuss the bargaining over wages and hours between workers and firms once a match has been formed. We will focus here on a bargaining scheme that leads to Pareto-efficient outcomes. Hours and wages are bargained over simultaneously and the bargained wage and hours maximise the joint surplus of the firm and the worker.⁶ The outcome of the bargaining process is obtained by maximisation of the

⁶Also, in the baseline we abstract from any wage rigidity. We will look at a different, right-to-manage bargaining scheme in section 4, and in section 5 we will analyse what real wage rigidity means for inflation and unemployment dynamics under the respective bargaining schemes.

Nash product:

$$\max_{w_t, h_t} (W_t - U_t)^\eta (J_t - V_t)^{1-\eta}. \quad (18)$$

The first term of the Nash product is the surplus a worker obtains from a job, raised to the worker's relative bargaining power $\eta \in (0, 1)$. The second term is the firm surplus, raised to the firm's relative bargaining power $1 - \eta$.

Under efficient bargaining, the first-order conditions can be shown to simplify to

$$\frac{\chi}{\lambda_t} h_t^\phi = x_t \alpha h_t^{\alpha-1} = \alpha \frac{x_t y_t^l}{h_t}. \quad (19)$$

This shows that the wage paid in any individual matches is not allocative for hours in existing matches, that is $\partial h_t / \partial w_t = 0$. Note that x_t , the price of the intermediate goods, coincides with the marginal cost for price-setting firms. Rearranging the above equation

$$x_t = \frac{\chi \frac{h_t^\phi}{\lambda_t}}{\alpha h_t^{\alpha-1}} \equiv \frac{mrs_t}{mpl_t} \quad (20)$$

highlights the implications of efficient bargaining for marginal costs (and inflation). Real marginal cost will equal the worker's marginal rate of substitution between consumption and leisure (mrs_t) divided by the real marginal product of labour (mpl_t). Both of these will, in turn, depend on hours worked per employee. Efficiency requires that the marginal gain to the firm of an additional hour worked (the firm's marginal value product of labour, $x_t mpl_t$) is the same as the marginal cost to the worker of that hour (his marginal rate of substitution between consumption and leisure, mrs_t). This makes clear that this subjective price of work determines the marginal wage, the marginal cost of producing intermediate goods and the marginal cost of producing differentiated goods. Note that, in contrast, the *average* hourly wage rate, w_t , does not measure marginal costs in the above setup.⁷

This bargaining procedure maximises the joint surplus of the match, that is $J_t + W_t - U_t - V_t$.⁸ The first-order condition for the wage bargaining simplifies to

$$\eta (J_t - V_t) = (1 - \eta) (W_t - U_t) \quad (21)$$

which states that each of the contracting parties receives a share of the total surplus

⁷The marginal cost here is determined by having an employed worker work for an additional hour, while the average wage rate averages over all hours worked (including the infra-marginal hours).

⁸Note that all terms in the wage cancel in the joint surplus of workers and firms since the wage is not allocative in efficient bargaining.

proportional to its relative bargaining power.

2.5 Vacancy posting

In order to find a worker, and so in order to produce, labour firms have to post a vacancy first. In the baseline we follow Mortensen and Pissarides (1999) and Merz (1995), and assume a constant, per period, cost of opening up a vacancy, $k_t^v = \kappa$.⁹ Free entry in the market for vacancies implies that in equilibrium the value of a vacancy is $V_t = 0$ in every period t and the equation for the asset value of a vacancy (15) can be recast as:

$$\kappa = q_t \mathbb{E}_t \{ \beta_{t,t+1} J_{t+1} \} . \quad (22)$$

where total vacancy costs are given by κv_t , and the average cost per hiring is κ/q_t . Using this, the value of a job for a firm simplifies to

$$J_t = y_t^l x_t - h_t w_t + (1 - \rho) \frac{\kappa}{q_t} . \quad (23)$$

From equations (22) and (23), one easily obtains the dynamic representation of the average cost per hire

$$\frac{\kappa}{q_t} = \mathbb{E}_t \left\{ \beta_{t,t+1} \left[y_{t+1}^l x_{t+1} - h_{t+1} w_{t+1} + (1 - \rho) \frac{\kappa}{q_{t+1}} \right] \right\} \quad (24)$$

which determines job creation in the model.

2.6 Monetary policy, fiscal policy and market clearing

Throughout the paper we assume that monetary policy is conducted according to the following Taylor rule

$$\hat{r}_t = \rho_r \hat{r}_{t-1} + (1 - \rho_r) \delta_\pi \hat{\pi}_{t-1} + \zeta_t^m \quad (25)$$

with $\rho_r \geq 0$ and $\delta_\pi > 1$ and ζ_t^m is an i.i.d. interest rate shock.¹⁰ Lump-sum taxes, t_t , are set so as to balance the budget period by period and there is no autonomous

⁹We look at alternative assumptions about vacancy posting costs in section 8.

¹⁰The Taylor rule that we use in this paper requires only lagged information. It is therefore implementable. A further advantage of using this specification is that the initial reaction of nominal rates to a monetary easing is the same in all the model variants that we discuss. We include only inflation and lagged rates in the Taylor rule. This ensures that the nominal rate is not directly influenced by differences in the output dynamics, which improves comparability across model variants.

government spending. Finally, equilibrium in the final goods market implies¹¹

$$y_t = c_t . \quad (26)$$

3 Stylised facts and the baseline economy

The aim of this paper is to elicit the role of specific labour market frictions for the behaviour of marginal cost and inflation. In order to obtain a measure of the empirical relevance of the respective approaches, we compare the responses to a monetary policy easing in the respective model variant to stylised facts for the euro area.

3.1 Calibration of the model economies

In order to put all variants on the same footing, the calibration is harmonised as much as possible. The calibration largely follows Christoffel *et al.* (2009) and matches salient features of the euro area between 1984 and 2006. Table 1 presents the assigned common parameter values.

The time-discount factor, β , is chosen to match an average annual real rate of 3.3%. We set ϕ to 10, implying an elasticity of substitution for labour of 0.1, in line with – but at the lower end of – microeconomic studies for the euro area.¹² The value of the risk aversion coefficient is set to $\sigma_c = 1.5$, following Smets and Wouters (2003).

Turning to the labour good sector and the labour market, we set $\alpha = 0.99$, so there are almost constant returns to hours per employee.¹³ On monthly data ranging to the early 1990s, Burda and Wyplosz (1994) estimate an elasticity of matches with respect to unemployment of $\vartheta = 0.7$ for France, Germany and Spain. Petrongolo and Pissarides (2001) survey estimates of the matching function for European countries and for the United States and conclude that a range from $\vartheta = 0.5$ to $\vartheta = 0.7$ is admissible. We select the midpoint, setting $\vartheta = 0.6$. The bargaining power of workers is set to a conventional value: $\eta = 0.5$.

For the variants of the model we consider later that involve wage stickiness, we follow the recent evidence of du Caju *et al.* (2008), which reports average wage contract

¹¹We assume that vacancy posting costs are taxes that are rebated lump-sum to the consumers and that do not require real resources. This simplifies the comparability of the different modelling variants.

¹²The elasticity of inter-temporal substitution of labour, $1/\phi$, is small in most microeconomic studies (between 0 and 0.5) for the euro area; for details, see Evers, Mooij, and Vuuren (2005), who report statistics based on a meta sample as well as estimates based on Dutch data.

¹³Our model does not contain capital. In some of the variants, there is thus a trade-off between obtaining a reasonable labour share (of around 60%) and small *ex post* profits associated with jobs. We opt to fit the latter and choose $\alpha = 0.99$ instead of a lower value.

<u>Preferences</u>		
β	.992	Time-discount factor (matches annual real rate of 3.3%).
ϕ	10	labour supply elasticity of 0.1.
σ_c	1.5	Risk aversion.
<u>Bargaining, Intermediate Producers and Labour Market</u>		
α	.99	labour elasticity of production (close to constant returns to scale).
ϑ	.6	Elasticity of matches with respect to unemployment.
η	.5	Bargaining power of workers
ξ_w	.83	Calvo stickiness wages (avg. wage duration 6 qtrs), where applicable.
<u>Monopolistic Retail Sector and Price Setting</u>		
μ_p	1/1.1	Price markup of 10%.
ξ_p	.75	Calvo stickiness of prices (average duration of 4 qtrs).
<u>Monetary Policy</u>		
γ_π	1.5	Response to inflation in Taylor rule.
γ_R	.85	Interest rate smoothing coefficient in Taylor rule.

Table 1: PARAMETERS AND THEIR CALIBRATED VALUES. The table reports calibrated parameter values that are identical over the different modeling approaches. The model is calibrated to the euro area from 1984Q1 to 2006Q4.

durations for various euro area countries of between one and three years. Where applicable, we therefore set the degree of nominal wage rigidity to $\xi_w = 0.83$, which implies an average wage duration of 6 quarters.

In the price-setting sector, we calibrate the markup to a conventional value of 10%, so $\mu_p = 1/1.1$ (in the Dixit-Stiglitz setup this implies a price-elasticity of demand of $\epsilon = 11$). For the average contract duration of prices we use the results of Alvarez *et al.* (2006) and set the corresponding Calvo parameter to $\xi_p = 0.75$, which amounts to an average price duration of 4 quarters. As is conventional, we set the response of monetary policy to inflation to $\gamma_\pi = 1.5$ and allow for interest rate smoothing by setting $\gamma_R = 0.85$.

Given these parameters, we set the other parameters of our models to ensure that the steady-state values for certain endogenous variables coincided across the different setups. The target values for these variables, shown in Table 2, thereby implicitly define the values of the remaining model parameters.

The steady-state unemployment rate is targeted to be $u = 9\%$, in line with the average of the euro area unemployment rate over the sample 1984 to 2006. We target a probability of finding a worker for an open vacancy of $q = 0.7$, in line with the euro area evidence collected in Christoffel *et al.* (2009). Our target for steady-state hours

	Value	Explanation
u	9%	Unemployment rate
q	70%	Probability of finding a worker
h	1	Hours worked per employee.
$\frac{b}{wh}$	65%	Unemployment benefits replacement rate.
ρ	.06	Quarterly separation rate.

Table 2: CALIBRATION TARGETS. The table reports calibrated parameter values. The model is calibrated to the euro area from 1984Q1 to 2006Q4.

worked per employee is $h = 1$. We set the replacement income b so as to ensure that the replacement rate in steady state in each of the variants, $\frac{b}{wh}$, equals 65%, in line with the net replacement rates published by the OECD in its set of ‘Benefits and Wages’ data. Finally, the evidence collected in Christoffel *et al.* (2009) and Hobijn and Sahin (2007) points to quarterly separation rates, from a worker flow perspective, of $\rho = 6\%$.

3.2 Stylised facts – the transmission of monetary policy

In response to a monetary shock that causes interest rates to fall by 100 basis points in annualised terms, empirical evidence from the literature and our own VAR (see Appendix) shows that in the euro area:¹⁴

Fact 1: output rises significantly above its steady state. Depending on the study the peak response varies between 0.3 and 0.7 percentage point.

Fact 2: inflation rises. The peak increase in year-on-year inflation lies between 0.2 and 0.3 percentage point.

Fact 3: wages per employee also rise but by less than output (in percentage terms).

Fact 4: employment rises significantly, and unemployment falls. Unemployment falls by 3% to 4.5% (number of heads), and the unemployment rate falls by 0.25 to 0.4 percentage point.

Fact 5: most of the adjustment in labour is borne by the number of employees rather than by hours worked per employee. Some uncertainty surrounds this statement,

¹⁴While significant differences exist in the microstructure of the economy, and the labour market in particular, between the euro area and the United States the above stylised facts happen to be consistent also with the evidence for the United States, e.g., Christiano *et al.* (2005), Angeloni *et al.* (2003) and Trigari (2005).

though, as the data for hours worked in the euro area are of relatively poor quality.¹⁵

In the following, we assess each of the different modeling variants against these five stylised facts.

3.3 Effects of a monetary policy shock in the baseline

As a first check against this evidence, Figure 1 reports the response of the economy to a monetary easing, where the impulse is an exogenous one percentage point reduction in the nominal interest rate. As can be seen in Figure 1, in the baseline inflation responds far too strongly to the monetary easing (by a factor of ten) while output effects are of the appropriate size. The strong response of inflation comes from the fact that real marginal costs for price setting firms are largely determined by the cost of increasing hours worked by each worker. Under efficient bargaining, the baseline case, these marginal costs are given by the marginal rate of substitution divided by the marginal product of labour, $x_t = \frac{mrs_t}{mpl_t}$. Given near constant returns to scale (α close to unity), the marginal product of labour, mpl_t , will be little affected. The percentage change in real marginal cost will therefore be driven by the response of the marginal rate of substitution, mrs_t . This in turn depends on the percentage change in the marginal utility of consumption and, fundamentally, on the response of hours per worker, where the latter is amplified ϕ times. Given our calibration, this implies a strong response of real marginal cost to an increase in hours worked and output.¹⁶

The remainder of the paper alters features of the model that affect the response of inflation to shocks. We start by assessing a different bargaining scheme.

4 Right-to-manage bargaining

In this section, we consider an alternative process for bargaining over wages and hours. In particular, we follow Trigari (2006) and use a ‘right-to-manage’ assumption for the

¹⁵For example, no quarterly series for actual hours worked exists for the euro area, or individually for all its member states. The measured response of hours worked (and hours per worker) therefore naturally is based on proxy series; cp. Appendix A.

¹⁶While the average wage rate does not have a direct bearing on inflation, there is an indirect effect. Employment and unemployment are determined by the job creation condition, (24). As illustrated in Figure 1, in the baseline model the real wage response to the monetary policy shock is large. Trigari (2006) highlights that solving the Shimer (2005) puzzle in the benchmark model also helps to reduce the response of inflation. If more of the labour adjustment needed to produce the additional output is provided through the extensive (number of employees) margin, the intensive (hours per employee) margin will also react less, which curbs the rise in the marginal rate of substitution and thus the rise in marginal costs.

bargaining process. In this setup, the wage rate is agreed upon first and then the firm is free to choose hours worked at that wage rate so as to maximise profits. The same wage rate applies to marginal and inframarginal hours worked. This setting seems reasonable for a number of wage contracts. In addition, it also preserves much of the empirically successful structure of the New Keynesian model.¹⁷ The first-order condition for labour demand is:

$$x_t \alpha h_t^{\alpha-1} = w_t \iff x_t = \frac{w_t h_t}{\alpha y_t^l} = \frac{w_t}{mpl_t}. \quad (27)$$

meaning that the cost of the marginal hour is equal to its benefits, i.e., that the hourly wage is equal to the marginal value product of an hour worked. The right-to-manage assumption radically modifies the composition of real marginal cost. Under efficient bargaining, real marginal cost depends on hours and the marginal utility of consumption. Under right-to-manage, the hourly wage becomes an essential element of real marginal cost, opening what Christoffel *et al.* (2009) and Trigari (2006), for example, term the ‘wage channel’. Therefore, the response of marginal cost and inflation to shocks will depend entirely upon the response of the bargained hourly wage to shocks.¹⁸

This notwithstanding, the effect of the right-to-manage assumption on marginal cost and inflation in equilibrium is *a priori* unclear. The major difference to the efficient bargaining assumption is that the choice of hours depends directly on the average wage, and that $\partial h_t / \partial w_t < 0$. This affects the wage bargaining. Wage equation (21) becomes

$$\eta \Omega_t (J_t - V_t) = (1 - \eta) (W_t - U_t) \quad (28)$$

with

$$\Omega_t = \frac{1}{1 - \alpha} \left(\frac{mrs_t}{x_t mpl_t} - \alpha \right) = \frac{1}{1 - \alpha} \left(\frac{\chi h_t^{1+\phi-\alpha}}{\alpha \lambda_t x_t} - \alpha \right) \quad (29)$$

This means that the relative bargaining power of the workers in the wage negotiation is modified by Ω_t : equations (28) and (29) suggest that the workers’ bargaining power increases with hours worked.¹⁹ The more inelastic the labour supply at the intensive

¹⁷Following Erceg *et al.* (2000), when modeling wage rigidities the literature typically assumes that monopolistically competitive labour firms produce an intermediate labour good. The assumption is that these firms set wages (the equivalent to the wage rate here) in a staggered manner and let labour output and thus input of hours into labour-good production (the equivalent to hours per employee here) be demand-determined.

¹⁸This does not, of course, mean that the marginal rate of substitution plays no role in inflation dynamics under right-to-manage. There is still an indirect influence of the marginal rate of substitution on inflation through its effect on the bargained wage.

¹⁹If we keep the free-entry assumption and constant recurrent vacancy posting costs, equation (22) still holds. The value of a job for a firm (23) and the dynamic of the average cost per hiring (24) can be simplified by replacing $x_t y_t^l$ by $h_t w_t / \alpha$.

margin (the larger is ϕ), the stronger this effect. If ϕ is large enough, the average wage will tend to be more responsive under right-to-manage than under efficient bargaining because of the strong relative bargaining power effect. It is less clear, however, whether this also implies that marginal costs are more responsive. Under right-to-manage, the marginal wage coincides with the average wage, w_t , which is more responsive than the average wage would be under efficient bargaining. Under efficient bargaining, however, the marginal wage, given by mrs_t , is already more responsive than the average wage.

Figure 1 compares the response of the economy to a monetary shock under both bargaining schemes. As the above logic suggests, the average hourly wage rate, w_t , responds more strongly (and in fact much too strongly in view of the stylised facts) to a monetary easing under right-to-manage. Still, the marginal wage rates (which determine the marginal cost of firms) show a very similar movement under both bargaining schemes, as can be seen in the similar response of marginal cost under the two schemes, leaving the response of inflation to a monetary policy shock essentially unchanged.²⁰ We conclude that the bargaining scheme in itself has little implication for inflation dynamics. As the next section shows, however, different bargaining schemes can imply fundamentally different effects of *wage rigidity* on inflation dynamics.

5 Wage rigidity

A cornerstone of the canonical New Keynesian framework is that average wage rates and their stickiness are instrumental for inflation dynamics; see, e.g., Christiano *et al.* (2005). This section of the paper discusses the role of wage stickiness for inflation dynamics once search and matching frictions in the labour market are accounted for. The related New Keynesian literature is Trigari (2006), Christoffel and Linzert (2005), Christoffel and Kuester (2008), and Krause and Lubik (2007).

In the model, the role of wage stickiness for inflation dynamics depends on the type of wage bargaining along which the model economy is organised (efficient vs. right-to-manage) as well as on whether wage stickiness affects only existing matches or also new matches (and thus hiring incentives of firms). We assume that in each period only a fraction $(1-\xi_w^o)$ of all existing wage contracts is renegotiated. New hires are paid

²⁰The different response in the average wage rate caused by the bargaining power effect does, however, imply differences in the response of unemployment. The average wage rate responds by more under right-to-manage and thus leaves relatively less of the increased revenue in the hands of the firms, thereby reducing hiring incentives. As a consequence, this leads to a smaller response of employment under right-to-manage than under efficient bargaining. This in turn leads to slightly more recourse to the intensive margin under right-to-manage than under efficient bargaining, which reinforces the increase in the wage and thus marginal costs.

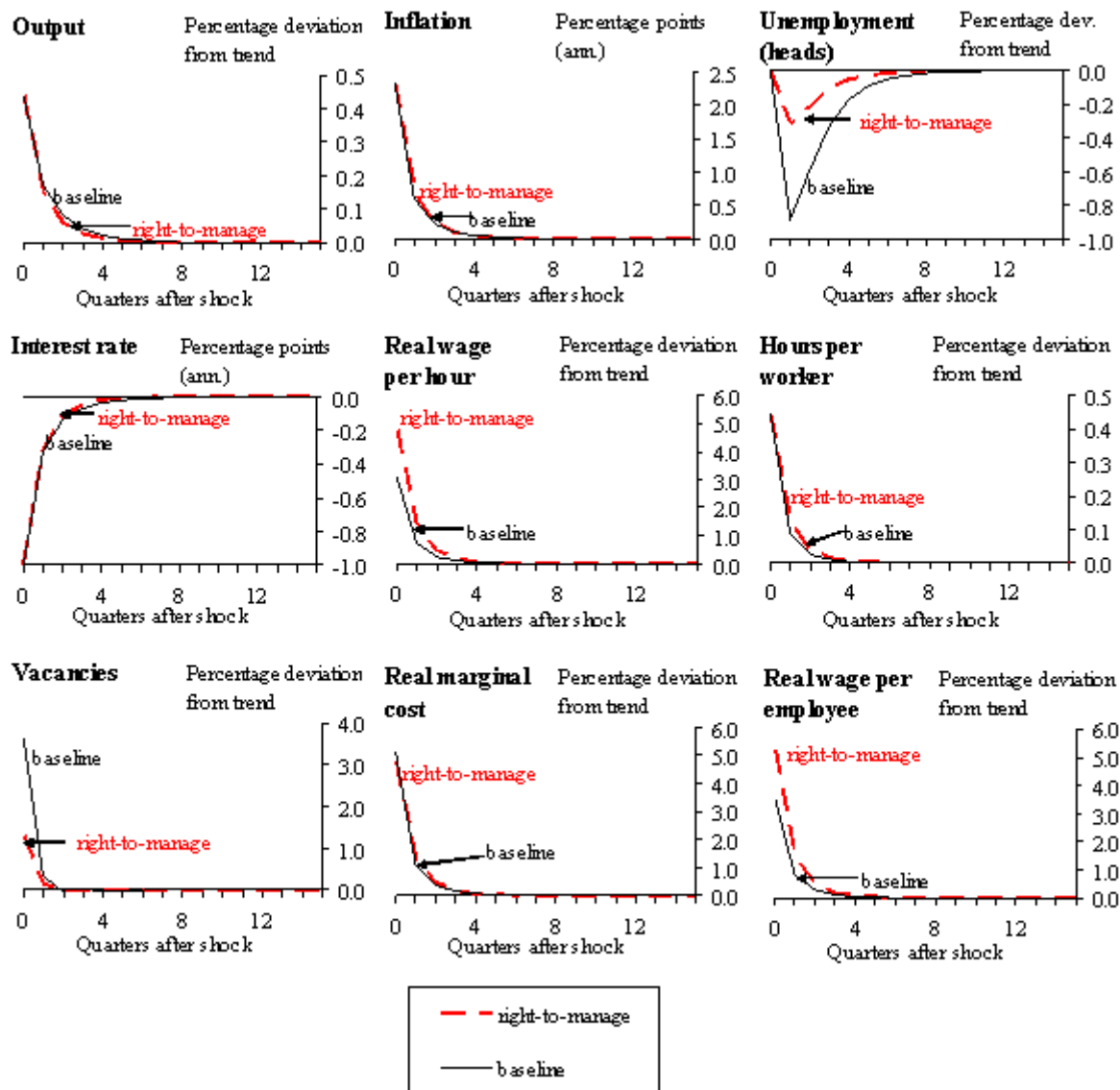


Figure 1: IMPULSE RESPONSES TO A MONETARY EASING – BASELINE VS. RIGHT-TO-MANAGE. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line marks the baseline response with efficient bargaining. The red dashed line shows the response under right-to-manage bargaining.

either the existing nominal contract wage from the previous period (with respective probabilities ξ_w^n) or they freely negotiate the wage (with probability $1 - \xi_w^n$). Full nominal wage flexibility obtains if $\xi_w^o = \xi_w^n = 0$.²¹

Even though the wage bargaining will be discussed in detail later, it is important at this stage to stress that all the ‘firm-worker’ pairs that are given the opportunity to (re)-negotiate their wage contract face the same problem and therefore set the same wage. Because wage negotiation is time-dependent, different workers may be paid different wages and can supply different hours, even though they are otherwise ex-ante identical.

5.1 Efficient bargaining

Under efficient bargaining, as we showed earlier, real marginal cost will equal the workers’ marginal rate of substitution divided by their marginal product and will be independent of the average wage rate in existing matches.²² For that reason, wage rigidity, if it only affects existing matches, does not have a bearing on inflation dynamics if bargaining is efficient.

Matters change if wage stickiness affects the wages of prospective new hires. As the vacancy posting condition equation makes clear, hiring incentives depend on the expected profits of firms. Wages are important in allocating the surplus of the match among firms and workers. As Hall (2005) notes, wage stickiness in new matches enhances the cyclical nature of job creation by altering the share of revenue left to firms over the cycle. Since hiring incentives are affected, future marginal costs and, thus, inflation will be affected as well. The reason is that the hiring behaviour of firms influences the economy-wide relative use of the extensive margin and the intensive margin of employment and this will have a direct bearing on marginal costs, as a result of the decreasing returns to hours worked per employee, $\alpha \leq 1$, and the increasing marginal disutility of work $\phi > 0$.²³ Wage stickiness, insofar as it affects new matches,

²¹These matches draw a wage from the previous period’s wage distribution. The rationale for this assumption, following Gertler and Trigari (2006), lies in interpreting a match as one position in a multiworker firm. While every period there may be new hires for some positions, these firms may adjust their overall pay scale only infrequently. There is currently debate about whether there is evidence in the data that wages of new hires are sticky. Pissarides (2007) and Haefke, Sonntag, and van Rens (2008), for example, find little empirical support for wage stickiness for new hires.

²²The reason is that in existing matches the average wage rate is not allocative but rather splits the surplus of the match among the two parties. In particular, even when fixing the *average* hourly wage rate, this does not fix the hourly wage rate *schedule*. As part of an efficient bargaining agreement, the relevant section of marginal wages can be freely set in a state-contingent way even if average wage rates are fixed.

²³See also Trigari (2006). Krause and Lubik (2007) find that wage stickiness is irrelevant for inflation dynamics. This rests on their assumption that there is only an extensive margin, in which case marginal costs indeed are little affected.

makes hiring more responsive to a monetary easing. This means that the demand-driven increase in labour input is borne more by the number of employees rather than by hours worked. As a result, the average worker works fewer hours than in the absence of wage stickiness. This means that the marginal rate of substitution rises less strongly and also that the marginal product of labour falls less strongly. So, wage stickiness for new hires, through its effect on employment, makes marginal costs (and thus inflation) less responsive to a monetary easing.

Figure 2 illustrates the empirical implications of this. In the baseline setup, equation (13) implies that employment reacts only with a lag. With wage stickiness in new matches, employment reacts more strongly. This leads to a quicker fall in the response of hours, lowering the persistence of the responses of marginal cost and inflation. While future marginal costs and inflation thereby react less than in the absence of wage rigidity, quantitatively the differences are tiny, as Figure 2 illustrates. Complementing the results in this section, Section 7 assesses the case when the baseline is modified to allow for contemporaneous hiring, in which case employment is a jump variable and the initial response of hours worked per employee is thereby attenuated.

5.2 Right-to-manage bargaining

Section 4 highlighted that right-to-manage implies a close relationship between hours worked and wages. In this section, we allow for stickiness in these wages. In particular, if the wage of a worker has been bargained i periods ago, he or she will work

$$h_t(w_{t-i}^*) = \left(\alpha \frac{x_t}{w_{t-i}^*} \frac{p_t}{p_{t-i}} \right)^{\frac{1}{1-\alpha}} = h_{t-i}(w_{t-i}^*) \left(\alpha \frac{x_t}{x_{t-i}} \frac{p_t}{p_{t-i}} \right)^{\frac{1}{1-\alpha}} \quad (30)$$

hours. Here, we let $w_{t-i}^* = \frac{w_{t-i}^{n,*}}{p_t}$ represent the real value of the nominal wage negotiated at time $t - i$. $w_{t-i}^{n,*}$ is the nominal wage rate negotiated i periods ago, which for some matches will continue to prevail today.

The asset value of a job clearly depends on the bargained wage. Adopting the viewpoint of an intermediate producer, we denote by $J_t(w_{t-j}^*)$ the asset value in period t of a job with a wage that was bargained i periods ago. *Ex ante*, the asset value of a new match for the firm can be written as

$$J_t = (1 - \xi_w^n) J_t(w_t^*) + \xi_w^n J_t(w_{t-1}) \quad (31)$$

where w_t is the average real value of all hourly wages. The asset value of a vacant job V_t is still given by equation (15); i.e., the vacancy posting condition is not modified.

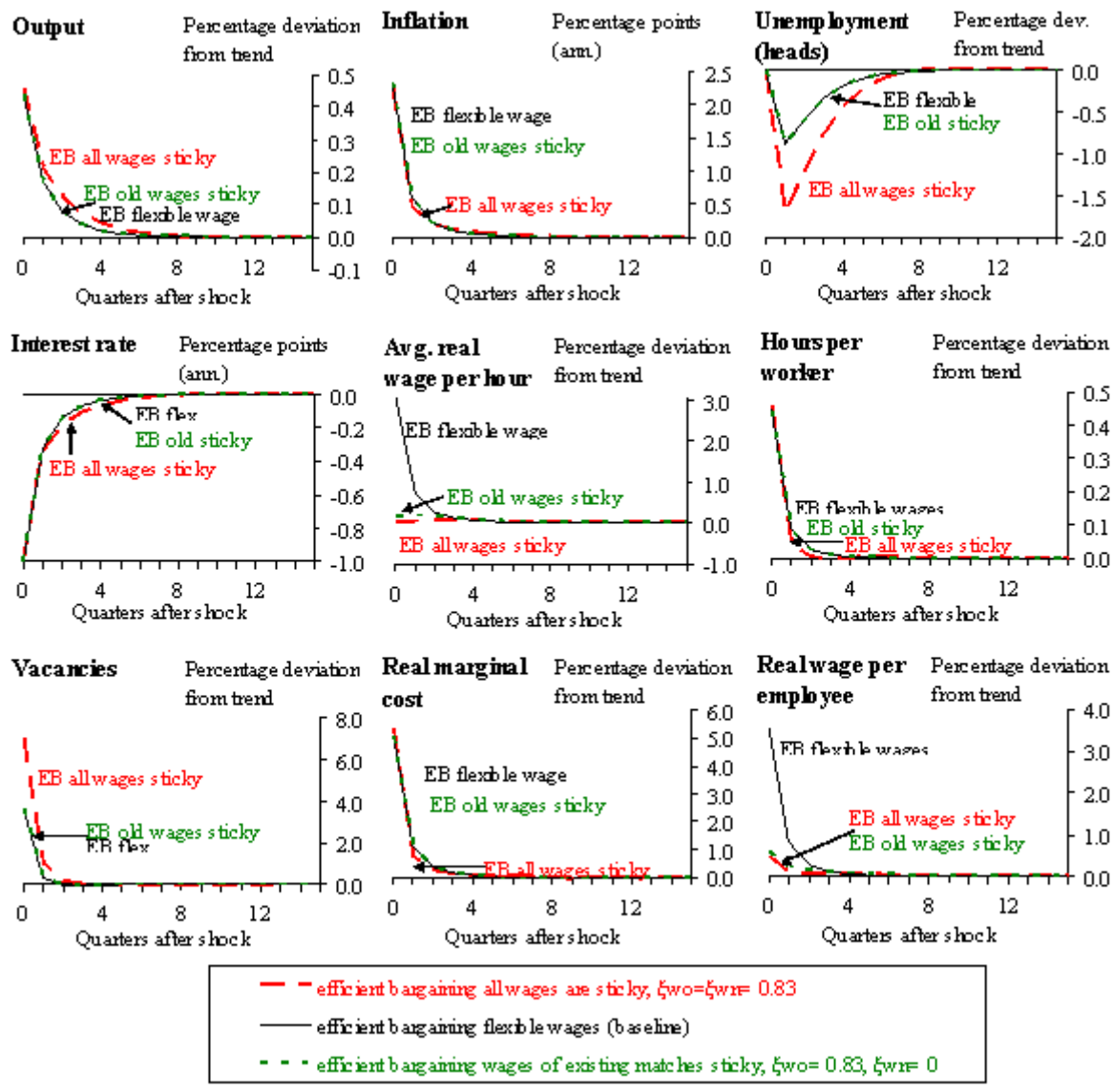


Figure 2: MONETARY EASING – EFFICIENT BARGAINING AND WAGE RIGIDITY OLD AND NEW HIRES. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line marks the baseline response. A red dashed line marks the response when both wages of new hires and of old matches are sticky. The green dotted line shows the response when only wages of old hires are sticky.

All the (re-)negotiating firm-worker couples face the same problem and therefore choose the same wage w_t^* through the usual Nash bargaining procedure

$$\max_{w_t^*} [\mathcal{W}_t(w_t^*) - U_t]^\eta [J_t(w_t^*)]^{1-\eta}$$

where the worker's surplus, $\mathcal{W}_t(w_t^*) - U_t$, takes into account the fact that wages may not be renegotiated immediately. After some algebra, we obtain the wage equation for wages negotiated at time t

$$\eta J(w_t^*) \Omega'_t = (1 - \eta) [\mathcal{W}_t(w_t^*) - U_t] \quad (32)$$

with

$$\Omega'_t = \frac{1}{1 - \alpha} \left[\frac{\chi h_t(w_t^*)^{1+\phi-\alpha}}{\alpha \lambda_t x_t} \frac{\mathbb{E}_t \sum_{i=0}^{\infty} (\beta(1 - \rho) \xi_w^o)^i \left[\frac{x_{t+i} p_{t+i}}{x_t p_t} \right]^{\frac{1+\phi}{1-\alpha}}}{\mathbb{E}_t \sum_{i=0}^{\infty} (\beta(1 - \rho) \xi_w^o)^i \left[\frac{x_{t+i}}{x_t} \left(\frac{p_{t+1}}{p_t} \right)^\alpha \right]^{\frac{1}{1-\alpha}}} - \alpha \right]$$

Figure 3 shows the response to a monetary easing in the right-to-manage model with flexible wages (which are similar to the efficient bargaining baseline) and the two alternatives with nominal wage rigidities. Two observations are in order. First and foremost, the combination of the right-to-manage bargaining with nominal wage stickiness produces much more rigid wages. These translate into a more reasonable response of inflation and into inflation persistence. In other words, the combination of sticky wages with 'right-to-manage' bargaining is able to generate a magnitude of the response of inflation to a monetary policy shock that is in line with the stylised facts. This follows directly from the fact that in right-to-manage bargaining, there is a direct cost channel from wages to marginal cost to inflation. Second, the responses are very similar for the two cases of wage rigidity. Whether wage rigidity affects only existing matches, or all matches, does not matter much for inflation dynamics; Christoffel and Kuester (2008) make a similar point.

6 Real rigidities arising within the individual firm

Following most of the existing literature, the baseline assumed that price-setting firms can buy labour goods, y_t^l , in a competitive factor market at cost x_t per unit. They then transform this into a differentiated product. Due to this assumption, price-setting firms' marginal costs are independent of their own output level. In that setup only if

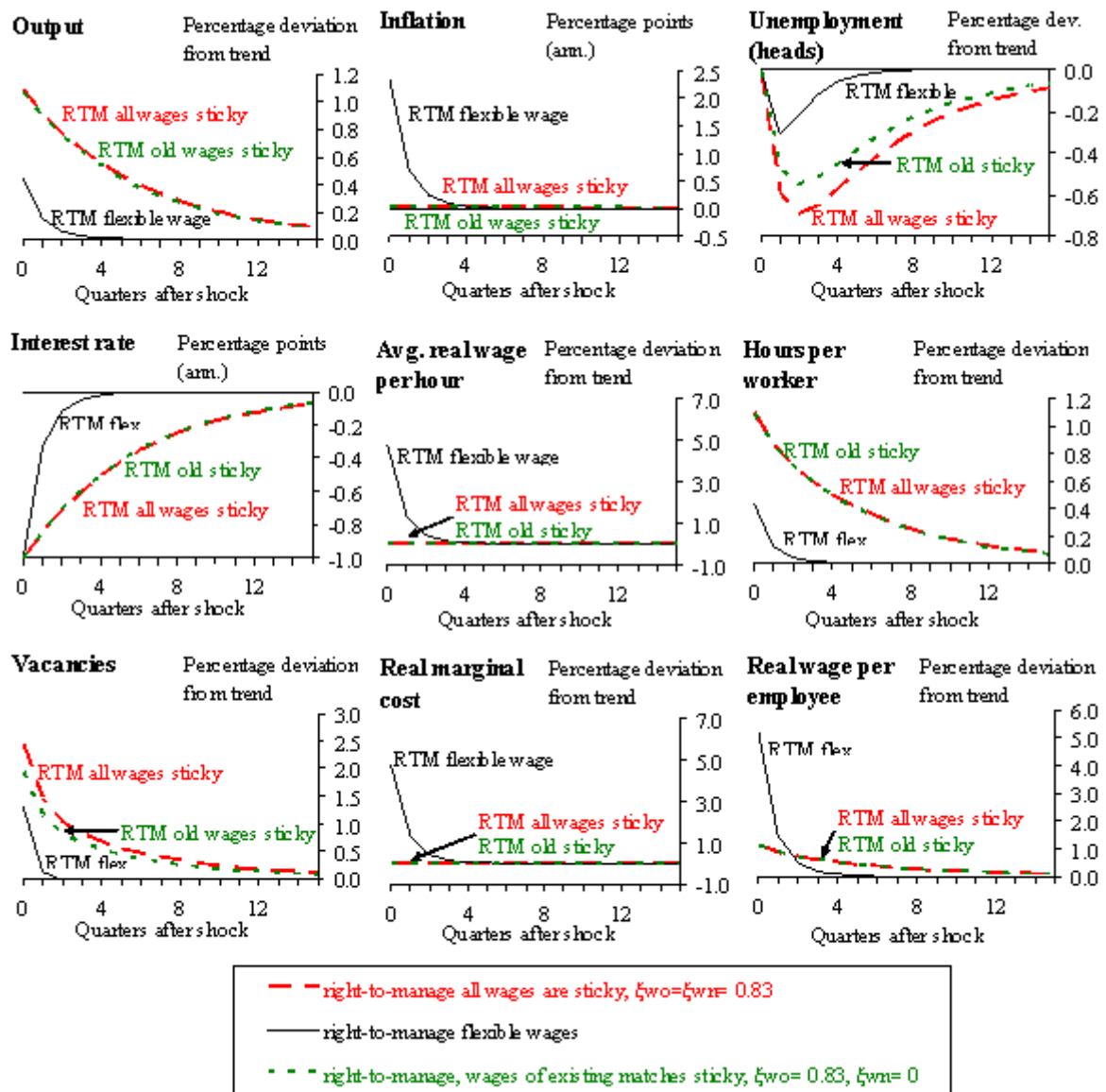


Figure 3: IMPULSE RESPONSES TO A MONETARY EASING RIGHT-TO-MANAGE WITH WAGE RIGIDITY. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line shows the response in the right-to-manage model in the absence of wage rigidity. The red dashed line shows the response when all wages are sticky. The green dotted line shows the case in which wage rigidity affects only existing matches.

aggregate marginal costs are more rigid will inflation be more rigid. Following Kuester (2007), this section instead merges the intermediate labour good sector with the retail sector and emphasises real rigidities arising at the *individual firm level* over and above the real rigidities arising at the aggregate level. Related papers are Sveen and Weinke (2007) and Thomas (2008).

The presence of search and matching frictions in the labour market means that a worker in this economy temporarily constitutes a firm-specific factor of production to the firm at which he is employed.²⁴ This, in turn, means that a firm's price setting has an effect not only on the demand that the firm faces but also on the wage demand of its worker and thus on the firm's own marginal costs.

As a consequence, for any given behaviour of aggregate marginal costs, firms are induced to adjust prices by less. The mechanism at work is the following. Consider an aggregate shock that, all else equal, would imply an increase in the marginal cost of all firms. A firm that can re-optimize its price passes part of the cost increase on to consumers. The increase in its price causes a fall in demand. This fall in demand will be the stronger the more price-elastic is demand; i.e., the larger is $\epsilon = \frac{1}{1-\mu_p}$. In turn hours worked at the firm fall (the more the production function exhibits decreasing returns to scale the more hours worked fall; i.e., the smaller is α). So, in sum, the larger ϵ is and the smaller α is the more hours worked fall in response to a price increase. If workers have an increasing marginal disutility of work, i.e., $\phi > 0$, this fall in hours worked leads to a reduction in the worker's marginal disutility of work. Therefore, at the time the firm decides what price it will set, it anticipates that the price increase will induce an effect that balances the original increase in marginal costs: workers take part of the original cost increase on themselves by accepting lower marginal wage rates. This implies that price-setting firms decide to move their prices by less for any given behaviour of aggregate marginal costs than in the baseline model, curbing the response of inflation to shocks relative to the benchmark model.²⁵

This intuition is captured in the New Keynesian Phillips curve for the modified model

²⁴In the intermediate labor sector of the baseline model, workers also constituted a temporarily firm-specific production factor. By assumption, however, this sector operated under perfect competition and flexible prices.

²⁵It seems important to distinguish between real rigidities arising at the aggregate level and real rigidities arising at the individual firm level. For real rigidities arising at the aggregate level, as in Ball and Romer (1990), prices (and thus inflation) will respond the less to shocks the less (the aggregate component of) marginal cost responds to these shocks. For real rigidities arising at the individual firm level, prices will respond the less to shocks the more (the firm-specific component of) any firm's marginal cost rises with demand. For a further exposition, see also Woodford (2003, Chapter 3).

economy:

$$\hat{\pi}_t = \beta(1 - \rho)E_t \{\hat{\pi}_{t+1}\} + \frac{1 - \xi_p}{\xi_p} [1 - \beta(1 - \rho)\xi_p] \left\{ \frac{1}{1 + \frac{\epsilon}{\alpha}[(1 - \alpha) + \phi]} \right\} \hat{x}_t. \quad (33)$$

With respect to the baseline New Keynesian Phillips curve, equation (9), there is an additional term dampening the pass-through of marginal cost on inflation (underlined). As the discussion above suggested, the more price-elastic demand is (the larger ϵ is), the more curved the marginal disutility of labour is (the larger ϕ) and the faster the returns to hours per employee decrease at the firm level (the smaller α), the less will firms adjust prices to aggregate shocks, and therefore the weaker will be the response of inflation to its aggregate driving forces.

Figure 4 compares the impulse response to a monetary easing in the baseline model to the model with firm-specific labour. The additional factor in the New Keynesian Phillips curve (33) is equal to $\left\{ \frac{1}{1 + \frac{\epsilon}{\alpha}[(1 - \alpha) + \phi]} \right\} = 0.009$. As a result, inflation reacts considerably less to the monetary shock when allowing for firm-specific labour, bringing the response more in line with the stylised facts. In turn, this implies that the monetary easing provides more stimulus, which translates into a response of output that is larger than in the baseline.

7 Contemporaneous hiring

In the variants of the search model discussed above there was a one-period lag before a worker hired in t takes on the job. In light of the empirical evidence on short-term fluctuations in employment, it is useful to look at a variant allowing for contemporaneous hiring.²⁶ This section follows Ravenna and Walsh (2008) and assumes the following timing. As before, at the beginning of each period some firms and workers randomly separate. Firms that are not matched with a worker can post vacancies. Contrary to the case in previous sections however, the successful matches become productive in the same period. So employment evolves according to $n_t = (1 - \rho)n_{t-1} + m_t$. This timing assumption implies the following flow values for the value of a job to the firm and the value of a vacancy

²⁶Data for the U.S. show that more than 60% of unemployment spells take less than 14 weeks; see “Labor Force Statistics from the Current Population Survey.” For the euro area the proportion of short-term unemployed (up to a 2-month unemployment spell) in total employment was around 15% in the period between 2005q1 and 2008q2, according to the OECD Labor Force Survey. Including spells of up to 5 months implies a proportion of roughly 30% of short-term unemployment in total unemployment.

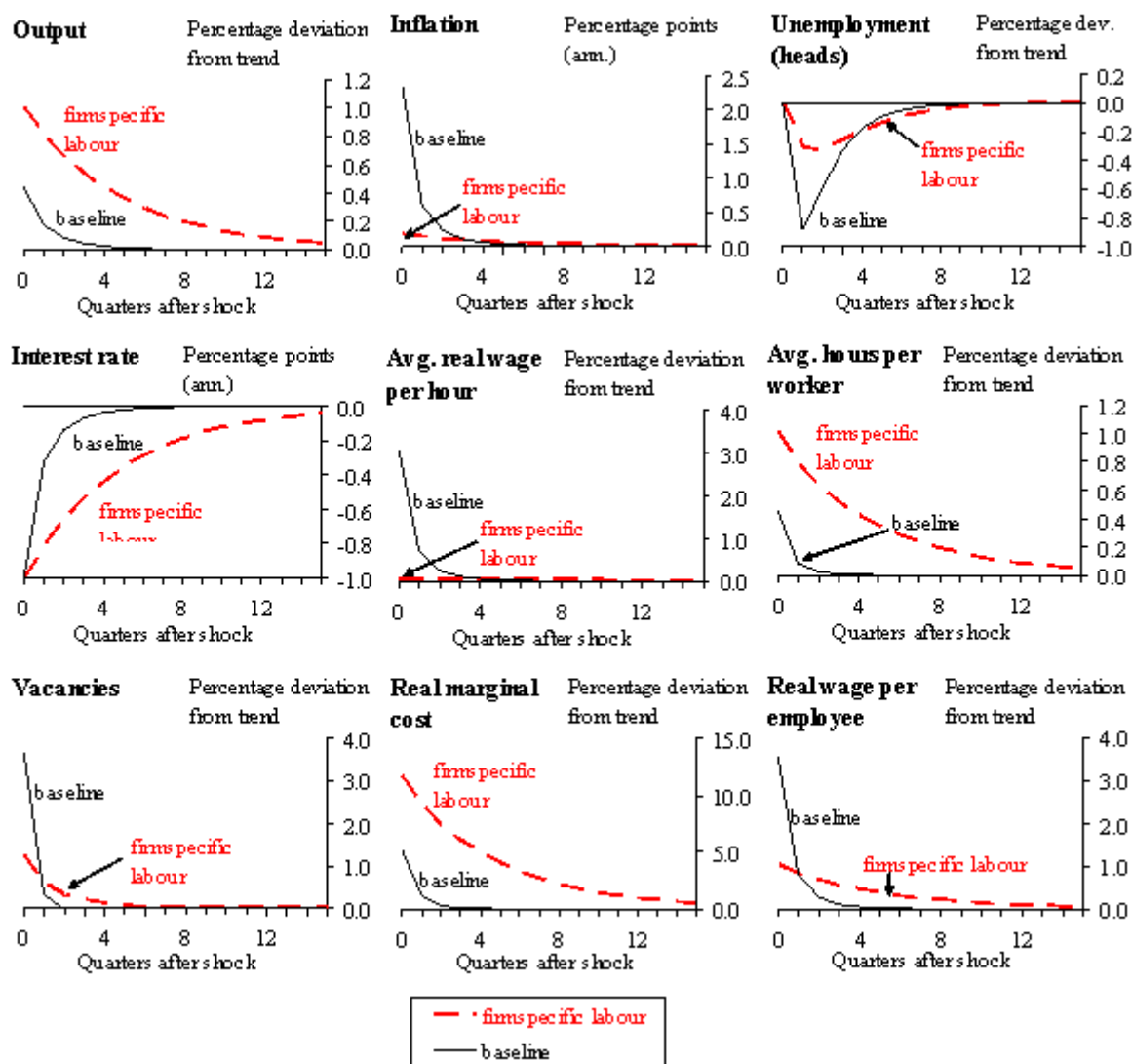


Figure 4: IMPULSE RESPONSES TO A MONETARY EASING – FIRM-SPECIFIC LABOUR VS BASELINE. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line shows the response in the efficient bargaining baseline, which separates price setting and wage setting into two different sectors. The red dashed line shows the response of the economy when these two sectors are merged.

$$J_t = y_t^l x_t - h_t w_t + \mathbf{E}_t \{ \beta_{t,t+1} (1 - \rho) J_{t+1} \} , \quad (34)$$

$$V_t = -\kappa + q_t J_t + \mathbf{E}_t \{ \beta_{t,t+1} (1 - q_t) V_{t+1} \} , \quad (35)$$

The flow value of being employed, W_t , and the flow value of being unemployed, U_t , are defined with respect to the end of the period after the formation of the current period matches.

$$W_t = h_t w_t - \frac{\chi}{\lambda_t} \frac{h_t^{1+\phi}}{1+\phi} + \mathbf{E}_t \{ \beta_{t,t+1} [(1 - \rho + \rho s_{t+1}) W_{t+1} + \rho(1 - s_{t+1}) U_{t+1}] \} , \quad (36)$$

$$U_t = b + \mathbf{E}_t \{ \beta_{t,t+1} [s_{t+1} W_{t+1} + (1 - s_{t+1}) U_{t+1}] \} . \quad (37)$$

Figure 5 shows the response to the monetary easing. The introduction of contemporaneous hiring has a direct impact on marginal costs. As matches become productive within the period, labour adjustment moves from the intensive to the extensive margin. As a result the reaction of employment is considerably more pronounced and the reaction of hours per worker is lower than in the baseline. The peak response of unemployment occurs in the first period and exceeds the peak response in the baseline by a factor of two. The strong response of employment implies a lower reaction of hours and a slightly lower wage response. The overall impact on the inflation response is small nevertheless.

8 Variants of the vacancy posting process

How vacancy creation moves over the business cycle is essential for determining the response of the intensive versus the extensive margin of labour adjustment. As discussed in section 3.3, this in turn has a bearing on the response of marginal wages, marginal cost and inflation over the business cycle. This section studies the implications of two modifications of the vacancy posting process that have been introduced in the RBC literature on frictional unemployment with the aim of helping the Mortensen and Pissarides model reproduce labour market stylised facts. The first modification, following Yashiv (2006), draws on the intuition that it is easier to incorporate a worker into a company when employment is high already. The second modification, following Fujita and Ramey (2007), eases the free-entry condition.

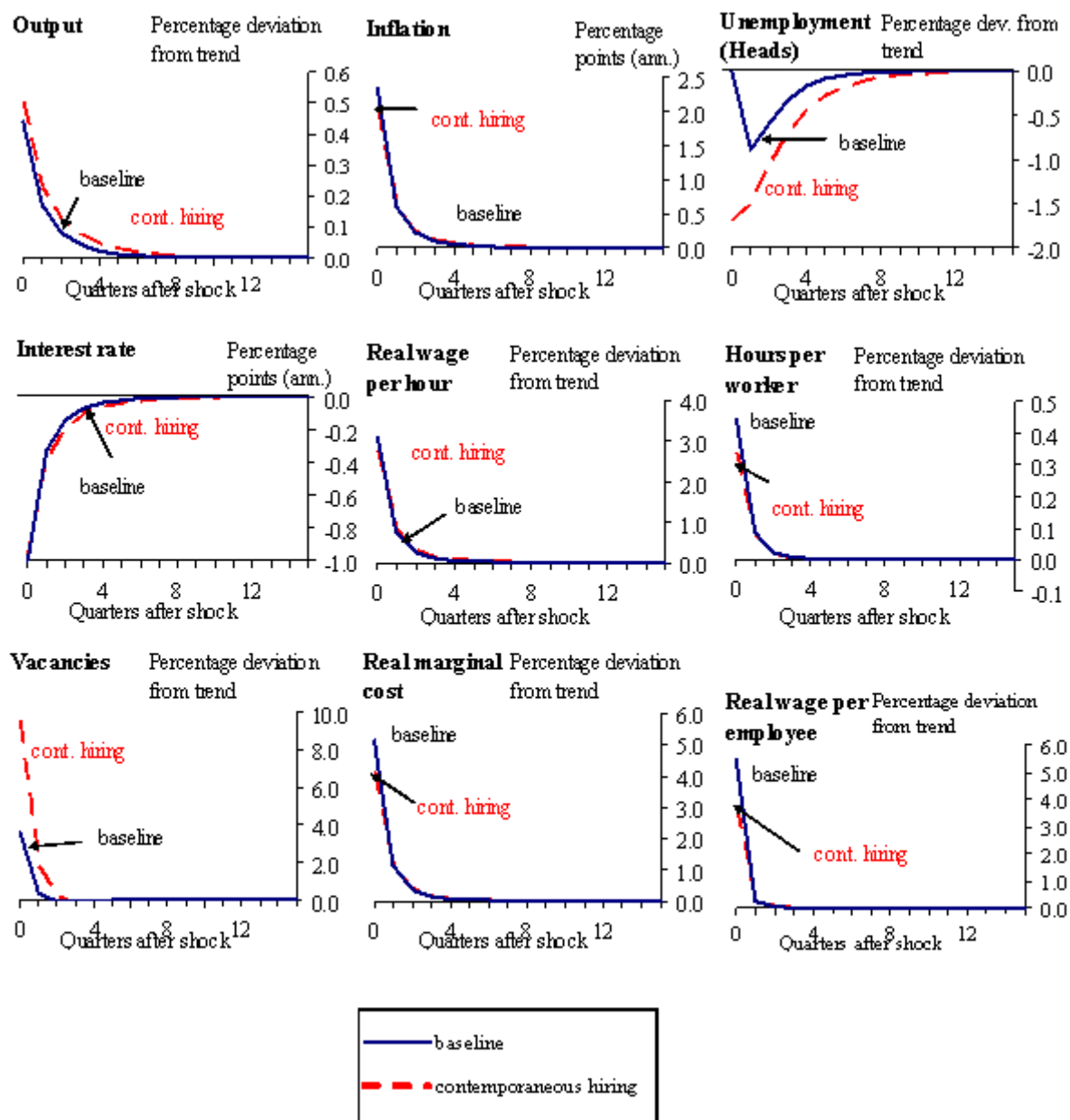


Figure 5: IMPULSE RESPONSES TO A MONETARY EASING – CONTEMPORANEOUS HIRING VS BASELINE. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line shows the response in the efficient bargaining baselines. The red dashed line shows the response of the economy when we allow for contemporaneous hiring.

8.1 Employment adjustment costs convex in the hiring rate

In the first variant of the model, we drop the assumption of a fixed recurrent vacancy posting cost. Instead, we follow Gertler and Trigari (2006) and assume that total hiring costs are a quadratic function of the hiring rate. The idea is that hiring costs not only increase with the number of new hires but also with the ratio of new hires to existing jobs,²⁷

$$k_t^v v_t = \psi \left(\frac{m_t}{n_t} \right)^2 n_t \quad (38)$$

where we assume that the total hiring cost is shared equally among all firms that posted a vacancy. Transforming (38), the cost per hiring, $k_t^v/q_t = \psi \frac{m_t}{n_t}$, is proportional to the hiring rate $\frac{m_t}{n_t}$. This is in contrast to the baseline where the cost per hire typically increases with n_t due to the congestion effect associated with the search and matching process. All else equal, this will lead to greater variation in employment and thus less variation in marginal costs and inflation. Using the free-entry condition $V_t = 0$ and substituting for k_t^v , equations (22), (23) and (24) are left unmodified but for the substitution of κ/q_t by $\psi m_t/n_t$, so that the dynamics of the job creation condition will be determined by the hiring rate instead of the probability of filling a vacancy.

Figure 6 shows that hiring reacts more strongly to shocks relative to the baseline, so future real marginal costs react less strongly, reducing the effect of the shock on inflation. Quantitatively, however, in our calibration the effect is tiny.

8.2 Vacancy costs as sunk costs

An alternative to convex costs is a sunk cost, as in Fujita and Ramey (2007). In this case, firms pay a sunk cost only once when they post a new vacancy. Every period, all the firms that have neither a worker nor a vacancy draw their sunk cost, K , out of a distribution, $F(K)$. Among these, all the firms endowed with a sunk cost lower than

$$\overline{SC}_t = \mathbb{E}_t \left\{ \beta_{t,t+1} \left[q_t J_{t+1} + (1 - q_t) \overline{SC}_{t+1} \right] \right\}$$

will post a vacancy.

²⁷The rationale for this way of modelling is in understanding a match as one job in large, multi-worker firms. The reason why costs per hiring decrease with the labour force is simply that more employment eases the integration of new hires.

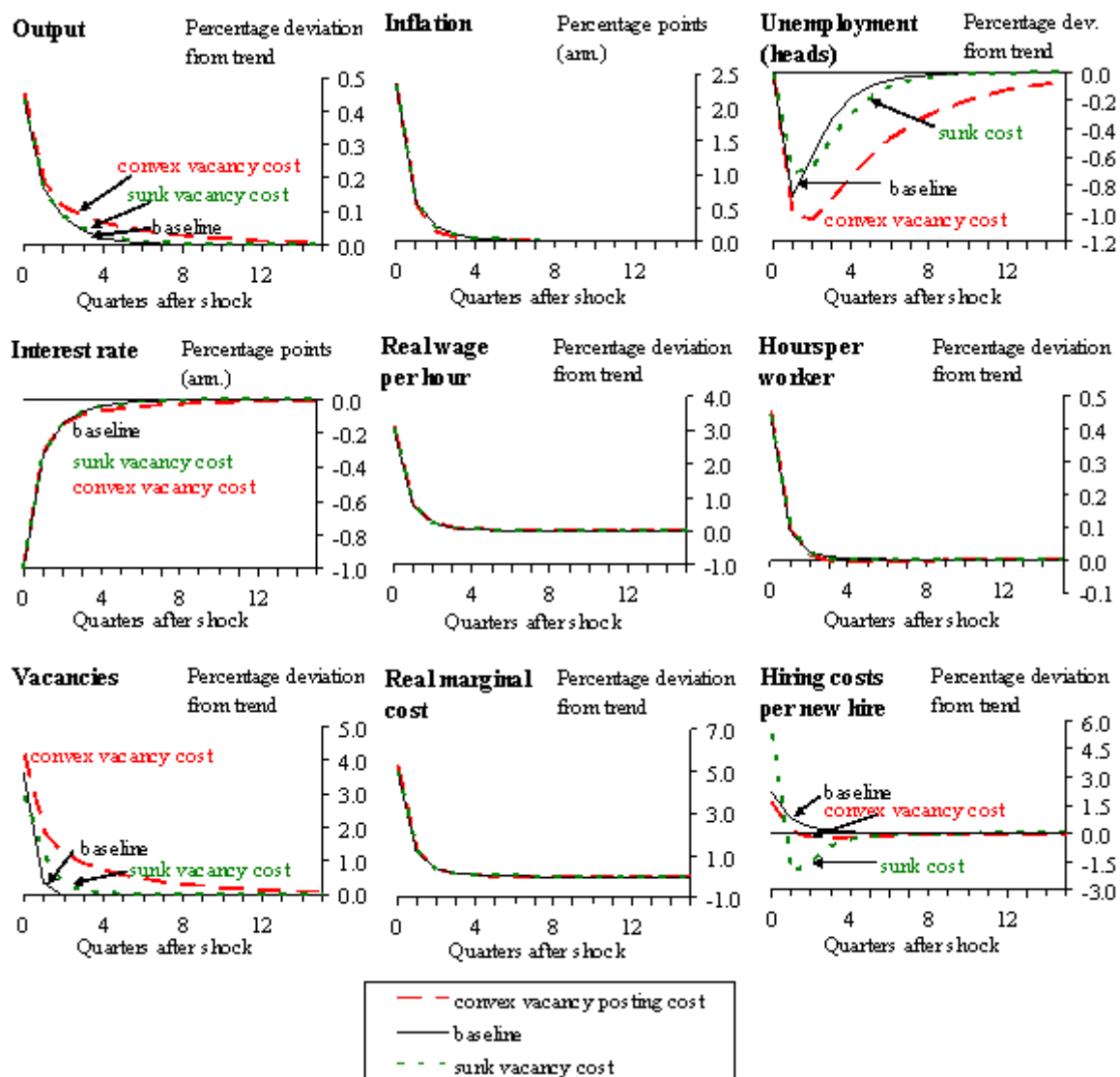


Figure 6: IMPULSE RESPONSES TO A MONETARY EASING WHEN ALLOWING FOR DIFFERENT VACANCY POSTING COSTS THAN IN THE BASELINE AND DIFFERENT VACANCY CREATION TECHNOLOGY. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The solid line shows the response in the baseline, which features constant costs per vacancy over the cycle, and in which these costs keep a vacancy open for exactly one period. The red dashed line represents the response in the variant with adjustment costs convex in the hiring rate. The green dotted line is the response in the variant where vacancy costs are sunk.

Denoting the newly posted vacancies by v_t^n , we obtain that

$$v_t^n = F(\overline{SC}_t)$$

and the law of motion of vacancies is:

$$v_t = (1 - q_{t-1}) v_{t-1} + v_t^n. \quad (39)$$

Let us assume that the cumulative distribution is a linear function of the sunk cost

$$F(K) = \varkappa(K - \bar{k}), \quad \varkappa > 0$$

where $\bar{k} \geq 0$ is the lowest sunk cost in the distribution while \varkappa is a slope parameter. Of course, the lower (respectively higher) is \bar{k} (respectively \varkappa), the higher is the number of new vacancies and the more volatile is employment. Finally, total vacancy posting costs are given by:

$$\varkappa \int_{\bar{k}}^{V_t} K dK = \varkappa \left(\frac{\overline{SC}_t^2 - \bar{k}^2}{2} \right) \quad (40)$$

Note that in this variant, no equation of the baseline model is modified except for that governing vacancy posting costs. The corresponding wage equation is obtained by substituting (14), (16) and (17) into (21) and cannot be further simplified. As illustrated by Figure 6 the addition of sunk costs clearly increases vacancy persistence (since vacancies are now a stock variable rather than a flow variable) but, again, the quantitative effects on the dynamic response of inflation to the shock are weak.

9 On-the-job search

In the baseline only unemployed workers can search, with the result that only they enter the matching function and can be matched up with vacancies. But, in reality, currently employed workers spend at least some of their time looking around for alternative job opportunities. The result of this is that job-to-job movements form a large part of all job destruction and creation and are, thus, potentially important for marginal costs over the cycle.²⁸

Following Krause and Lubik (2006) and van Zandweghe (2009), we assume that jobs

²⁸Akerlof *et al.* (1988) and Gomes (2008) suggest that job-to-job flows account for around 50% of all job separations in the United States and the United Kingdom, respectively. However, Contini and Revelli (1997) suggest that the proportion is smaller than this in euro area countries.

come in two types: ‘good’ jobs (that pay higher wages) and ‘bad’ jobs (that pay lower wages). Aggregate output is produced using workers employed in both types of jobs. The presence of these two types of jobs creates an incentive for workers employed in bad jobs to search for good jobs. So, good firms are able to recruit from a pool of workers that includes both those unemployed workers who are searching for good jobs, $u_{g,t}$, and those currently employed in bad jobs, $e_t = \zeta_t n_{b,t}$, where ζ_t is the search intensity of employed workers and $n_{b,t}$ is employment in the bad sector. Given this, the relevant measure of labour market tightness for the ‘good’ sector will be $\tilde{\theta}_t = \frac{v_t}{u_{g,t} + e_t}$, where $u_{g,t}$ is unemployment of good sector workers. This measure of market tightness will likely be less volatile than the measure in the baseline model without search on-the-job, which was given by $\theta_t = \frac{v_t}{u_t}$. The reason is that job-to-job flows, e_t , will be procyclical. In a boom, more new vacancies will be posted than if there were no on-the-job search; since employment in the ‘bad’ sector is procyclical, the number of searching workers will also be procyclical, increasing the likelihood of filling a vacancy in a boom relative to the model with no on-the-job search. The first-order condition for employed search intensity will be given by:

$$\tau\gamma\zeta_t^{\tau-1} = \frac{\eta}{1-\eta} (1-\rho) \sigma_{m,g} \left(\frac{v_{g,t}}{u_{g,t} + e_t} \right)^{1-\vartheta} \cdot \left(\frac{\kappa_g}{\sigma_{m,g}} \left(\frac{v_{g,t}}{u_{g,t} + e_t} \right)^{\vartheta} - \frac{\kappa_b}{\sigma_{m,b}} \left(\frac{v_{b,t}}{u_{b,t}} \right)^{\vartheta} \right) \quad (41)$$

where ζ_t^{τ} is the cost of searching with intensity ζ_t . As can be seen, as more vacancies are posted, employed workers increase their search intensity, which will lower expected hiring costs and increase the incentive to post vacancies even further. That is, the rising search activity of the employed in a boom forms an additional resource in the matching function that helps to keep hiring costs more stable. The model predicts that job-to-job flows rise in booms, in line with the data.

In the model with on-the-job search (and in the baseline model), the New Keynesian Phillips curve will continue to be given by (9). Given that final output is assumed to be a Cobb-Douglas function of output in each of the two sectors, real marginal cost will be given by:

$$x_t = \left(\frac{x_{g,t}}{1-\omega} \right)^{1-\omega} \left(\frac{x_{b,t}}{\omega} \right)^{\omega} \quad (42)$$

where ω denotes the share of the output of ‘bad’ jobs in final output. With efficient bargaining over hours, real marginal cost in each sector will again be equal to the

marginal rate of substitution divided by the marginal product of labour in that sector:

$$x_{g,t} = \frac{\chi h_{g,t}^{1+\phi}}{\lambda_t \alpha y_{g,t}^l} \quad (43)$$

and

$$x_{b,t} = \frac{\chi h_{b,t}^{1+\phi}}{\lambda_t \alpha y_{b,t}^l} \quad (44)$$

So the response of real marginal cost to demand shocks will depend upon the response of hours in each sector to such shocks. On-the-job search leads to a reduced response in hiring costs in response to shocks; as this affects the extensive margin, this will reduce the response of future real marginal cost to shocks and, thus presumably, the response of inflation.

Despite this, however, Figure 7 shows that even with on-the-job search, the response of inflation is only slightly smaller than the inflation response in the baseline model. In turn, this is because of the large and immediate 6% rise in real marginal cost in response to this shock, offset only slightly by the lower response of real marginal cost in future periods relative to the baseline.

10 Endogenous job destruction

In much of the literature on frictional labour markets job separation is treated as an exogenous process, for several reasons: it makes the model easier to solve; it makes the model easier to calibrate; and it implies a robust negative correlation between unemployment and vacancies, whereas models that endogenise separation often have problems replicating the Beveridge curve. However, the separation rate rises in recessions.²⁹ Therefore, in this section we consider a version of the baseline model with endogenous separation. This imposes an additional first-order condition related to the separation margin and thus further restricts the model's predicted dynamics.

As in Mortensen and Pissarides (1994), we endogenise separation by assuming that each match has an idiosyncratic productivity component that is uncorrelated with aggregate productivity. Thus, the income produced by a match is given by $x_t z_{it} h_{it}^\alpha$, where z_{it} is the match-specific productivity component and the rest of the notation follows that in

²⁹Sources of evidence for countercyclical separation in Europe include Burda and Wyplosz (1994). For the U.S., see Fujita and Ramey (2008). Shimer (2005) and Hall (2005) illustrate that this countercyclicality may be less pronounced than was previously believed.

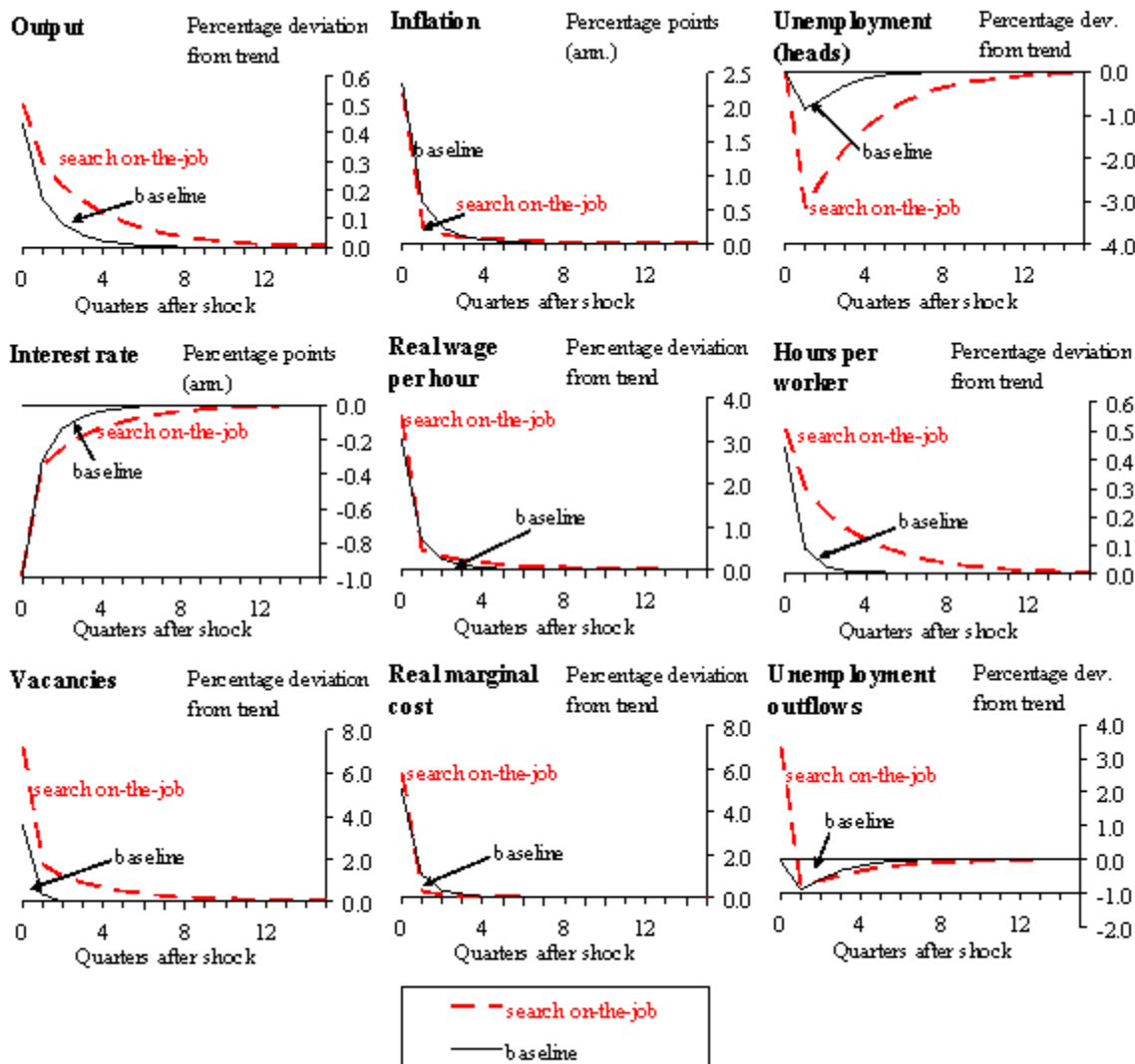


Figure 7: IMPULSE RESPONSES TO A MONETARY EASING WHEN ALLOWING FOR SEARCH ON-THE-JOB. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line shows the response in the baseline, which features homogeneous jobs, and in which workers are recruited only out of the unemployment pool. The red dashed line shows the response when allowing for search on-the-job.

Section 2. We will assume z_{it} is normally distributed with a cumulative distribution F , with mean 1 and standard deviation σ_z . This distribution will allow us to nest the exogenous search model as a special case of our endogenous search model associated with $\sigma_z \rightarrow 0$.³⁰ We will call the reservation productivity at which separation occurs \underline{z}_t .

Following den Haan, Ramey, and Watson (2000), we assume that z_{it} is *iid*. That is, z_{it} is uncorrelated with the aggregate shock x_t , it is uncorrelated across individual matches i , and most important, it is uncorrelated over time. For continuity with the rest of the paper, we still include an exogenous separation component: at the end of each period, existing matches separate for exogenous reasons with probability ρ^x . Allowing for productivity shocks with distribution F in all periods, and reservation productivity \underline{z}_t , the employment dynamics are:

$$n_t = (1 - F(\underline{z}_t)) [(1 - \rho^x)n_{t-1} + m_{t-1}] \quad (45)$$

The fraction of jobs surviving between periods $t - 1$ and t is therefore $1 - \rho_t \equiv (1 - F(\underline{z}_t))(1 - \rho^x)$. The total number of worker-firm pairs that meet at the end of $t - 1$ is $m_{t-1} = \sigma_m v_{t-1}^\vartheta u_{t-1}^{1-\vartheta}$, where $u_{t-1} = 1 - n_{t-1}$ but the number that actually enter into production at the start of t is only $(1 - F(\underline{z}_t))m_{t-1}$.

The rest of the model has the structure spelled out in Section 2. In particular, we assume efficient bargaining, as in Section 2.4, which now implies that both wages and hours will depend on the match-specific productivity shock z_{it} . Thus in the analogue to equation (14), a matched firm's value function is a function of the shock, $J_t(z_{it})$, and its flow of revenues is $x_t z_{it} h_t(z_{it})^\alpha - w_t(z_{it}) h_t(z_{it})$. A matched worker's value function $W_t(z_{it})$ also depends on the shock; the worker's flow payoffs are $w_t(z_{it}) h_t(z_{it}) - \lambda_t^{-1} \frac{\chi}{1+\phi} h_t(z_{it})^{1+\phi}$. On the right-hand side of all the Bellman equations there will therefore be expectations of future values with respect to z_{it+1} .

Efficient bargaining over hours gives a first-order condition that is equivalent to (20), except that it also takes into account the idiosyncratic productivity shock:

$$h_t(z) = (\alpha x_t \lambda_t / \chi)^{\frac{1}{1+\phi-\alpha}} z^{\frac{1}{1+\phi-\alpha}} \equiv H_t z^{\frac{1}{1+\phi-\alpha}} \quad (46)$$

Thus hours are proportional to $z^{\frac{1}{1+\phi-\alpha}}$ with the time-varying factor of proportionality H_t , which will help us aggregate hours and wages. This allows us to integrate

³⁰We will choose parameters such that continuation is optimal for $z = 1$, implying a reservation productivity strictly less than one. In the limit as $\sigma_z \rightarrow 0$, the fraction of separations occurring endogenously goes to zero and the mean of z for continuing matches goes to one. Therefore in this limit the additional first-order condition serves only to define the separation threshold, which has no effect on the remaining equations of the model.

across productivity to express all quantities in aggregate terms. After some algebraic manipulation, the job creation Euler equation analogous to (24) becomes

$$\frac{\kappa}{q_t} = E_t \left\{ \beta_{t,t+1} \left[\frac{(1-\eta)(1+\phi-\alpha)}{1+\phi} \frac{x_{t+1}c_{t+1}}{n_{t+1}} - (1 - F(z_{t+1})) (\eta\kappa\theta_{t+1} + (1-\eta)b) + (1 - \rho_{t+1}) \frac{\kappa}{q_{t+1}} \right] \right\} \quad (47)$$

The reservation productivity \underline{z}_t satisfies the job destruction equation $J_t(\underline{z}_t) = 0$. Eliminating the wage and expected future values from the Bellman equation for J , the job destruction equation can be written as

$$\frac{(1-\eta)(1+\phi-\alpha)}{1+\phi} x_t H_t^\alpha \underline{z}_t^{\frac{1+\phi}{1+\phi-\alpha}} - \eta\kappa\theta_t - (1-\eta)b + (1-\rho^x) \frac{\kappa}{q_t} = 0 \quad (48)$$

Calibrating this version of the model requires us to determine what fraction of separations are endogenous. For the U.S., Davis and Haltiwanger roughly calculated the fraction of endogenous separations in U.S. data by comparing firm-level job destruction (which they argue should reflect changes in firms' labour demand, and therefore map into the model's endogenous component) with total worker separation flows (which are larger and should include worker-initiated changes in employment status, which they map into the model's exogenous component). They find that the endogenous component is roughly 40% of the total, which – absent evidence for the euro area – is the number we will use here. Then for any standard deviation σ_z of the idiosyncratic shock, we can calculate $\underline{z} = F^{-1}(0.4 \cdot 0.06)$, assuming 6% total separation flows as before, and from \underline{z} we can calculate all the integrals over z that appear in the equations.

Other parameters can be backed out as before: σ_m can be deduced from observed job flows, and χ can be chosen so that (46) is satisfied. Finally, we can choose κ and b for consistency with the job creation and job destruction equations, (47) and (48). For a given σ_z , this makes the replacement ratio endogenous, so this calibration strategy will not automatically generate the same replacement rate used in previous sections of the paper. Therefore, we have simulated the model for a range of values σ_z , looking for a parameterisation that implies the same replacement ratio assumed in the baseline model. When we set the standard deviation of the idiosyncratic shock to $\sigma_z = 0.4$, we obtain a 65% replacement ratio, as in the baseline model.

For this parameterisation, Fig. 8 shows the impulse responses for endogenous separation together with those of the exogenous separation model that has the same replacement ratio (65%). The monetary stimulus primarily goes into wages. The overall stimulative effect of the money shock is relatively small.³¹ As the figure shows, hours

³¹ Interestingly, there are various parameter regions implying major qualitative changes in the

per worker and vacancies both rise, but separation and unemployment increase, too, for this calibration. While inflation and most other aggregate variables behave similarly in the two cases, the introduction of endogenous separation thus leads to counterfactual implications for unemployment. We thus confirm a previous finding in the literature, e.g., Krause and Lubik (2007), that even when endogenising separation has little impact on quantities outside of the labour market, it can be very hard to force all the margins of labour market adjustment to co-move realistically.

impulse responses, including regions of indeterminacy. For example, setting $\sigma_z = 0.1$, which leads to a replacement ratio of 84%, we find that a monetary stimulus causes a large fall in the separation rate. Under these parameters, unemployment is more volatile (and responds in the right direction) and wages and inflation are much smoother. Thus, while the parameterisation with $\sigma_z = 0.4$ fails to improve the model's inflation predictions, the effects of endogenous separation differ (and improve relative to the data) at $\sigma_z = 0.1$.

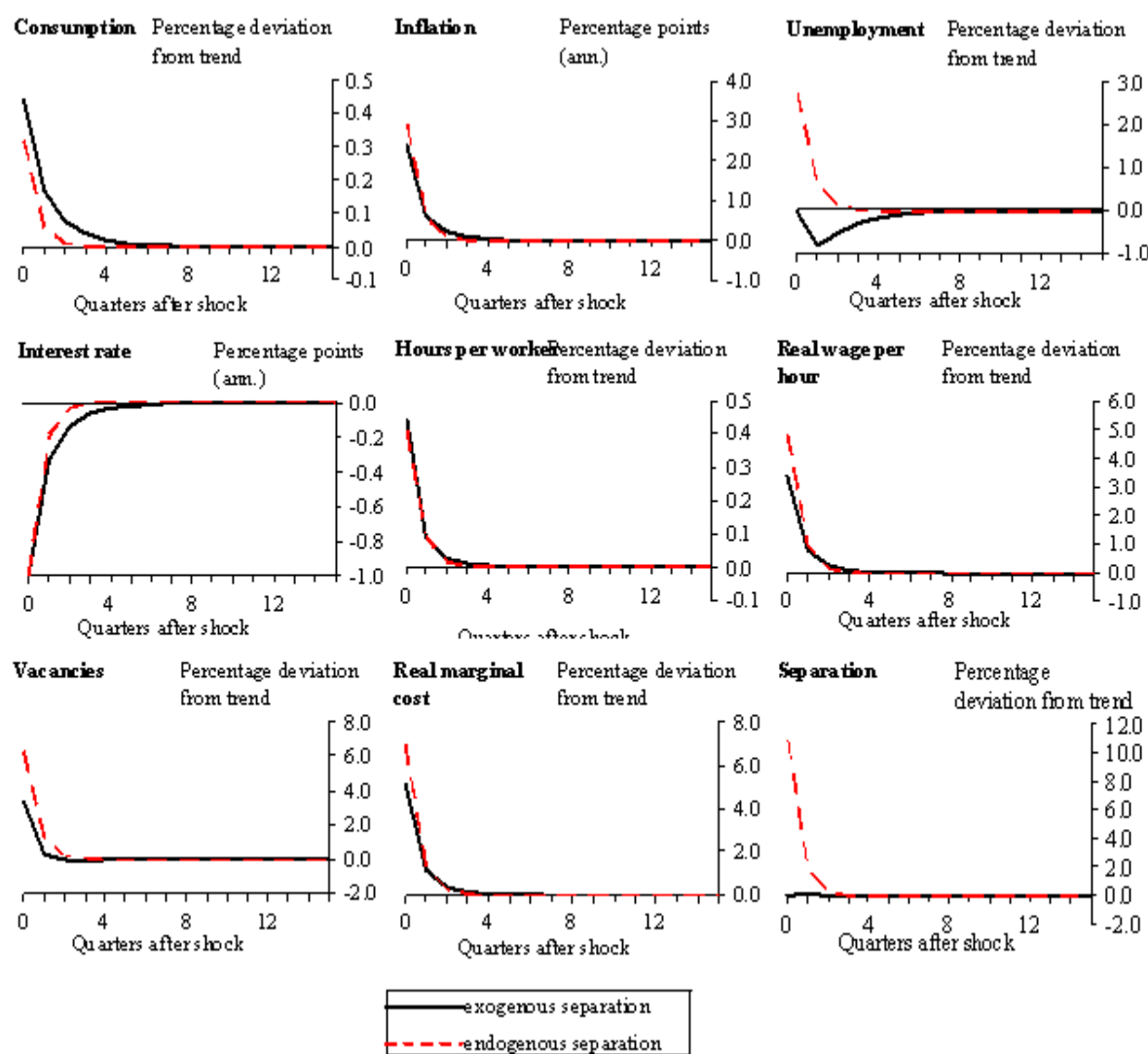


Figure 8: IMPULSE RESPONSES TO A MONETARY EASING, MODEL WITH ENDOGENOUS SEPARATION VS. MODEL WITH EXOGENOUS SEPARATION, ALTERNATIVE CALIBRATION. Shown are impulse responses to an unanticipated 100 bps reduction in the quarterly nominal interest rate in quarter 0. All entries are in percentage deviations from steady state. Interest rates and inflation rates are reported in annualized terms. The black solid line refers to the response in a model, which features only exogenous separation between firms and workers. Apart from slight differences in the calibration (described in the main text body) this resembles the baseline model. The red dashed line marks the responses when allowing for endogenous separation. Shown is the case with a low standard deviation of the idiosyncratic technology shock, $\sigma_z = 0.4$.

11 Conclusions

In this paper, we have assessed the ability of labour market variants in a New Keynesian model to explain inflation dynamics using, as our empirical benchmark, the responses of variables to a monetary policy shock in the euro area. Our results are largely negative and are collected in Table 3 in this section. We first used a standard New Keynesian model with search and matching frictions and efficient bargaining to show that such models typically generate a response of inflation to nominal shocks that is much too large relative to the data. We then considered many, if not most, of the alternative specifications that currently exist in the literature.

	Out-put	Infla-tion	Unem-ploy-ment	Em-ploy-ment	Real wages per empl.	Real wage per hour	Total hours wor-ked	Hours per empl-oyee
Euro area data	0.6	0.3	-4.5	0.4	0.3	0.4	0.9	0.04
Baseline (eff. barg.)	0.4	2.3	-0.9	0.1	3.5	3.1	0.4	0.4
RTM bargaining	0.4	2.3	-0.3	0.0	5.2	4.8	0.4	0.4
Sticky wages (EB)	0.5	2.3	-1.7	0.2	0.6	0.0	0.5	0.5
Sticky wages (RTM)	1.1	0.0	-0.6	0.1	1.1	0.0	1.1	1.1
Firm-level real rigidity	1.01	0.1	-0.3	0.03	1.03	0.03	1.02	1.02
Contempor. hiring	0.5	2.0	-1.7	0.2	3.1	2.8	0.5	0.3
Hiring costs (convex)	0.5	2.4	-1.1	0.1	3.6	3.1	0.5	0.5
Hiring costs (sunk)	0.4	2.4	-0.7	0.1	3.4	3.0	0.4	0.4
On-the-job search	0.5	2.2	-3.2	0.3	4.1	3.6	0.6	0.5
Endog. separation	0.3	2.9	2.8	-0.3	5.2	4.8	0.1	0.4

Table 3: PEAK RESPONSES TO A 100 BPS MONETARY EASING IN THE MODELS AND IN THE DATA. Top: euro area data (cp. the appendix for details). Then: baseline model with efficient bargaining (EB) and right-to-manage bargaining (RTM); models with EB and RTM and sticky wages for new and existing matches; real rigidities arising at the firm-level due to firm-specific labour; contemporaneous hiring instead of a one-period lag; different forms of hiring costs (period by period but convex in the hiring rate, or sunk); on-the-job search; and allowing for endogenous separation. Relative to the baseline, each of the models changes one feature at a time. From left to right: output, inflation, unemployment, employment, real wages per employee, real wage per hour, total hours worked, hours per employee. All variables are expressed in log percentage deviation from steady state.

We first considered the addition of wage rigidity. Under efficient bargaining, wage stickiness in existing jobs has no effect, since average wages in existing jobs are non-allocative. But wage stickiness in newly created jobs does have an effect. In this case, the existing wage rate will affect real marginal cost, since the wage paid to new hires

affects hiring incentives and, thus, marginal costs. Combined with a ‘right-to-manage’ assumption for the determination of hours (which implies a direct channel from wages to inflation), staggered wages at the level of the match will help to smooth the reaction of the aggregate wage, resulting in a smaller response to shocks of marginal cost and inflation. When we then combined this nominal rigidity with the real rigidity of firm-specific labour, we found that prices respond less to shocks the more any individual firm’s marginal cost responds to a change in that firm’s price. Hence, the response of inflation to nominal shocks is reduced if the marginal disutility of work is increasing in hours worked and/or there are decreasing returns to labour; both these conditions hold in our calibrated model.

Given the central role of employment dynamics for marginal costs, we then looked at variants of the hiring cost function that have been proposed to imply stronger and more persistent employment fluctuations. We found that neither of the modifications we considered materially affected the response of inflation to nominal shocks. Finally, we considered two other margins along which adjustment can occur in the labour market: on-the-job search and endogenous job destruction. We found that both features could help reduce the response of inflation to shocks. However, the effects were not particularly strong. In addition, the effects of endogenising job destruction depend critically on the calibration of the model, and some of the implications for other model variables were counterfactual.

In summary, we found that a model with sticky nominal wages when combined with right-to-manage bargaining was best able to capture the response of inflation to nominal shocks. Reassuringly, in that setup it was also largely irrelevant whether wage stickiness affected only existing matches or also new matches. More generally, our results suggest that the extent to which the labour market matters for inflation dynamics depends crucially on its microeconomic structure.

For future work, it would therefore be interesting to test the mechanisms we have explored using microeconomic data. Work on wage stickiness for new hires that has recently been conducted, e.g., by Haefke *et al.* (2008) is a first step in that direction. Microeconomic evidence on bargaining, and contractual arrangements regarding state contingency of the wage, overtime payments and premia, overtime accounts and non-wage components of compensation would be a valuable further step toward assessing the relative merits of these different labour market setups.

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12 Appendix A: further evidence on monetary transmission in the euro area

In the paper, we benchmarked the different approaches to modelling the labour market against five stylised facts about the transmission of monetary policy in the euro area. This section presents the background evidence from which we deduced these stylised facts.³² Figure 9 prints the impulse responses of euro area aggregates to a 100-basis-point monetary shock (an easing) in McCallum and Smets (2008), who estimate a factor-augmented VAR following the methodology of Bernanke *et al.* (2005).³³ Also shown are the boundaries of bootstrapped 90% confidence intervals. A monetary easing stimulates the economy. According to this evidence, at its peak output has risen by about 0.8% above steady state. Inflation increases, too, but in a muted manner. The peak increase in quarter-on-quarter GDP deflator inflation is about 25 basis points (in annualised terms). As a consequence of higher demand, the number of unemployed people, i.e., unemployment, falls by 3%.³⁴ The increase in economic

³²A number of studies have examined the response of real and nominal variables in the euro area to a monetary shock using alternative identification mechanisms. For early evidence on the euro area as a whole see, e.g., Peersman and Smets (2003) and Angeloni *et al.* (2003). More recent papers are McCallum and Smets (2008) and Peersman and Straub (2007). For evidence country by country, see, e.g., Mojon and Peersman (2003) and Normandin (2006).

³³Their data comprise quarterly euro area aggregates and a number of euro-area country-specific series from 1987Q1 to 2005Q4, as well as foreign series.

³⁴The depicted fall in unemployment corresponds to a fall in the unemployment rate of about 0.27 percentage point. Unemployment is calculated in a model-consistent way from the response in employment, which McCallum and Smets report, by assuming that the labor force does not react to a monetary easing and that average unemployment rates are 9%.

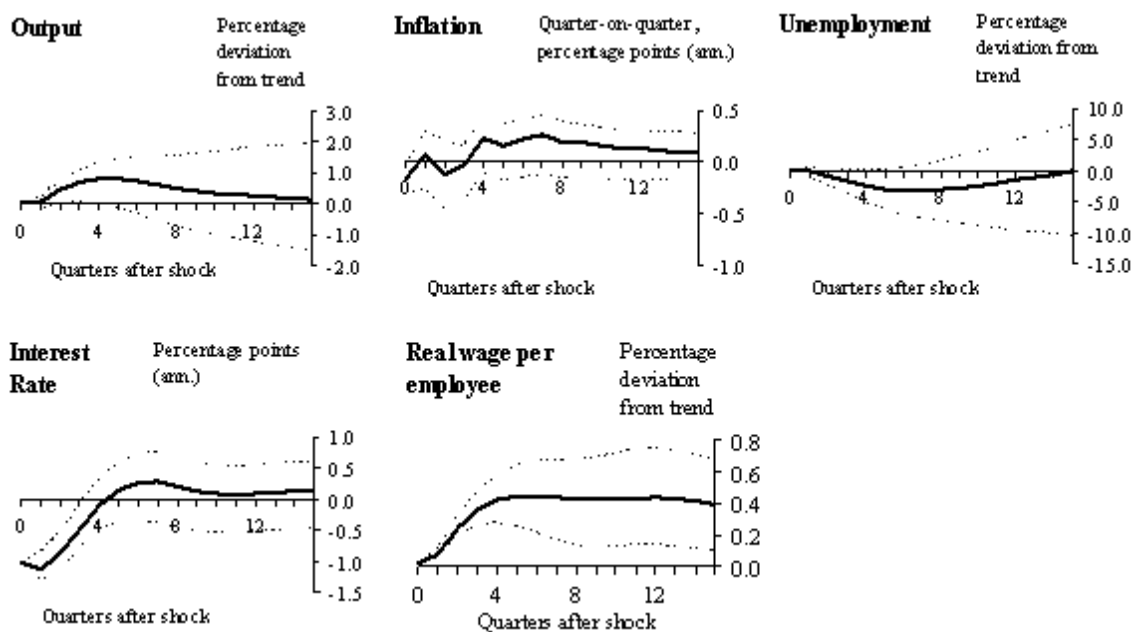


Figure 9: IMPULSE RESPONSES TO AN UNANTICIPATED MONETARY SHOCK IN THE EURO AREA, SOURCE: MCCALLUM AND SMETS (2008). Depicted are responses to a 100 bps monetary easing (annualized nominal rates). Charts refer to the euro area. The sample is 1987Q1 to 2005Q4. Top row (from left to right): response of real GDP (in percentage deviation from trend), quarter-on-quarter GDP inflation (in annualized percentage points), unemployment (in percentage deviation from trend). Note: shown is the response of log unemployment (heads), $\log(ut)$. Bottom row: response of nominal interest rate (in annualized percentage points) and the real wage per employee (in percentage deviation from trend). Reprinted with the authors' permission.

activity also translates into an increase in the real wage per employee. That increase is about half the size of the increase in output but is highly persistent.

Complementing this evidence, Figure 10 reports our own estimates of the monetary transmission in the euro area. The sample starts in 1987Q1, as in McCallum and Smets (2008), and runs through 2006Q4. We estimated a standard VAR(4) in the variables shown in the graphs and identified the monetary shock through the oft-used recursive ordering scheme. For details see the notes to Figure 10. According to these estimates, an unanticipated monetary easing induces a longer-lived deviation of nominal rates from the baseline than in McCallum and Smets (2008) and, as a result, most responses are more persistent. With regard to amplitudes, the evidence above is robust. A monetary easing of 100 basis points causes a significant increase in output of about 0.8% at its peak, and an increase in real wages per employee of about half that size (which is, again, highly persistent). Unemployment, according to our estimates, falls by 6%.³⁵ At the same time, the increase in inflation remains contained, with

³⁵This translates into a reduction in the unemployment rate of about 0.54 percentage points.

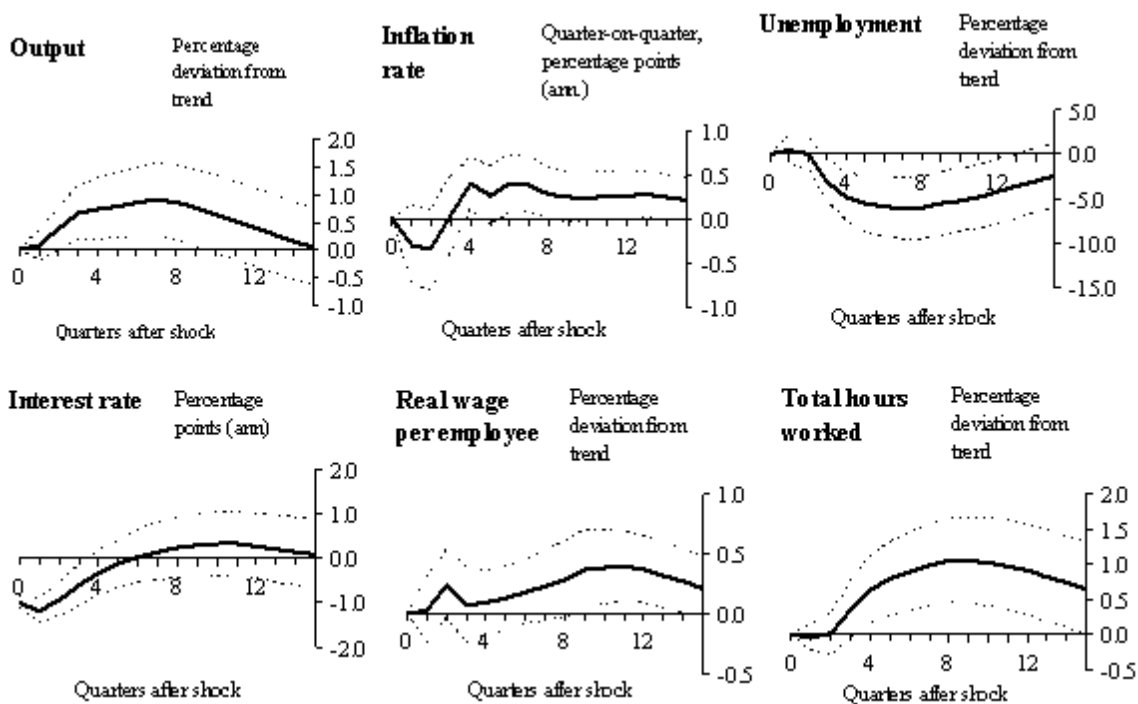


Figure 10: IMPULSE RESPONSES TO AN UNANTICIPATED MONETARY SHOCK – OWN ESTIMATES. Underlying the results is a VAR(4) estimated on above data (taken from the AWM data set except for hours worked, which is taken from Christoffel et al., 2009). The sample is 1987Q1 to 2006Q4; identification relies on the usual Cholesky decomposition. The data were de-trended variable by variable prior to running the VAR. For real variables, we allow for a trend, and a break in the trend plus a level shift in 1991Q1 so as to accommodate German reunification. Inflation and nominal interest rates were regressed on a linear trend, which breaks after 1998Q4 in order to account for the disinflation in the euro area prior to the introduction of the common currency, and the constancy of the inflation objective thereafter. Dotted lines mark asymptotic 90% confidence intervals (5% and 95% bounds).

GDP deflator inflation remaining below 0.4 percentage point in annualised terms. In the above estimation, we resort to the same proxy for hours worked in the euro area as used by Christoffel *et al.* (2009).³⁶ When using this series to infer hourly wage rates and hours worked per employee, the results suggest that changes at the extensive margin are more relevant for labour adjustment than hours worked per employee. The first two panels of Figure 11 suggest that hours worked per employee react little throughout the period (and if anything, they fall), while the number of employees is more responsive. As a result, the increase in wages per employee depicted in Figure 10 is mainly attributed to a slow but persistent increase in hourly wages instead of an

³⁶Unfortunately, harmonized quarterly series for hours worked do not exist for the euro area. The proxy uses the quarterly real GDP series for the euro area to interpolate annual EU-Klems data on total hours worked in the euro area.

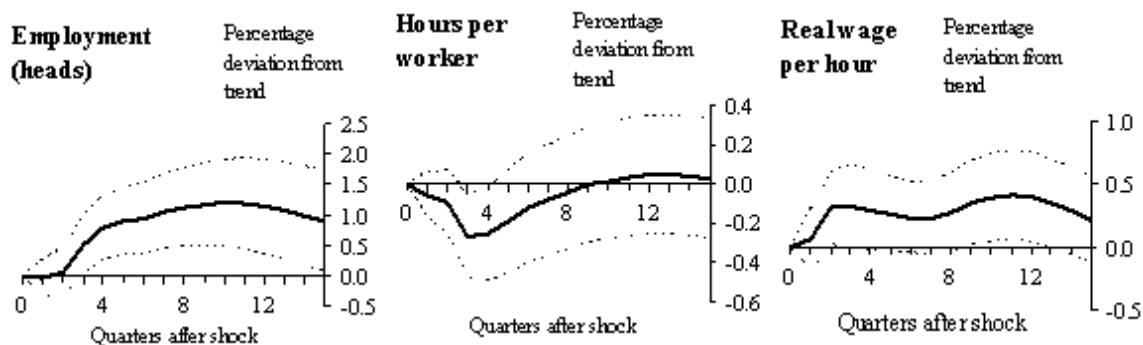


Figure 11: IMPULSE RESPONSES OF HOURS AND WAGES TO A MONETARY SHOCK – OWN ESTIMATES. Shown are the responses of employment, hours per employee and the real wage rate. The VAR is the same as in Figure 10, except that the measures of hours per worker and of the real hourly wage rate replace the measures for total hours worked and the real wage per employee.

increase in hours per worker. (See Figure 11.) Due to the limitations of the euro area data with respect to hours worked or hourly wages, however, the responses in Figure 11 – while indicative – should be taken with some caution.

Finally, Table 4 summarises the impulse responses reported above and augments these with the evidence of three further studies. Angeloni *et al.* (2003) report impulse responses for the euro area based on various identification schemes for the monetary shock in a VAR. So do Peersman and Smets (2003). The table, in addition, collects evidence of a recent study by Peersman and Straub (2007), who identify a monetary shock (and other shocks) based on sign restrictions, which were derived from a prototypical New Keynesian model. The table summarises the maximum amplitudes of the variables following the monetary shock. Most of the studies imply that after the initial shock, nominal rates fall somewhat further before being tightened. In order to harmonise the studies, all responses have been normalised by the maximum easing of the nominal interest rate that these studies report.

	Out-put	Infl.	Un-empl.	Empl.	Real wages per empl.	Real wage per hour	Total hours worked	Hours per empl.
McCallum,Smets (1)	0.7	0.2	-2.9	0.3	0.4	–	–	–
Own VAR (2)	0.6	0.3	-4.5	0.4	0.3	0.4	0.9	0.04
Peersman,Straub (3)	0.3	0.2	-2.8	0.3	0.3	0.3	0.3	0
Angeloni et al. (4)	0.5	0.2	-3.3	0.3	–	–	–	–
Peersman,Smets (5)	0.5	0.2	-3.3	0.3	–	–	–	–

Table 4: PEAK RESPONSES TO A MONETARY EASING IN THE LITERATURE. Most of the studies imply that after the initial easing, nominal rates fall somewhat further before they tighten. In order to harmonize the studies, all responses have been normalized by the maximum easing of the nominal interest rate that these studies report (which in all studies occurs one quarter after the shock). Entries in small type refer to numbers that have been deduced indirectly as described below. (1): McCallum and Smets (2008) report the response of employment. The response for unemployment is inferred from the figure for employment assuming a constant labor force and an average unemployment rate of 9%. (2): Own VAR: hours worked data for the euro area are subject to measurement errors as stated in the main text. (3): Peersman and Straub (2007) only report the response of total hours worked but state that responses are similar when employment is used. We thus report the same response of employment as for total hours worked. The response for unemployment is inferred as in (1). Similarly, hourly wage rates and hours per employee are inferred from this statement. (4): Peersman and Smets (2003): sample 1980-1998. Entries are based on their Fig. 2.1 and Fig. 2.8. The response for unemployment is inferred as in note (1). (5): Angeloni *et al.* (2003) report the response of the unemployment rate (see their Table 6). The response of unemployment reported here is obtained using an unemployment rate of 9%. Figures for employment were obtained accordingly. Results for the Angeloni *et al.* paper are averages over different specifications for the sample 1980-2001. Only the response of consumer prices is reported. Results were read off the figures and tables in their paper.

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